How must I conduct statistical comparisons in my Experimental Study? On the use of Nonparametric Tests and Case Studies.





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In this talk

We focus on the use of statistical tests for analyzing the results obtained in a design of experiments within the fields of Data Mining and Computational Intelligence.

Motivation

The experimental analysis on the performance of a new method is a crucial and necessary task to carry out in a research on Data Mining, Computational Intelligence techniques.

Deciding when an algorithm is better than other one may not be a trivial task.

Motivation

Deciding when an algorithm is better than other one may not be a trivial task.

Example for classification

Large Variations in Accuracies of Different Classifiers

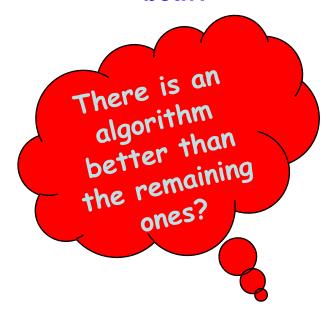
	Alg. 1	Alg. 2	Alg. 3	Alg. 4	Alg. 5	Alg. 6	Alg. 7
aud	25.3	76.0	68.4	69.6	79.0	81.2	57.7
aus	55.5	81.9	85.4	77.5	85.2	83.3	85.7
bal	45.0	76.2	87.2	90.4	78.5	81.9	79.8
bpa	58.0	63.5	60.6	54.3	65.8	65.8	68.2
bps	51.6	83.2	82.8	78.6	80.1	79.0	83.3
bre	65.5	96.0	96.7	96.0	95.4	95.3	96.0
cmc	42.7	44.4	46.8	50.6	52.1	49.8	52.3
gls	34.6	66.3	66.4	47.6	65.8	69.0	72.6
h-c	54.5	77.4	83.2	83.6	73.6	77.9	79.9
hep	79.3	79.9	80.8	83.2	78.9	80.0	83.2
irs	33.3	95.3	95.3	94.7	95.3	95.3	94.7
krk	52.2	89.4	94.9	87.0	98.3	98.4	98.6
lab	65.4	81.1	92.1	95.2	73.3	73.9	75.4
led	10.5	62.4	75.0	74.9	74.9	75.1	74.8
lym	55.0	83.3	83.6	85.6	77.0	71.5	79.0
mmg	56.0	63.0	65.3	64.7	64.8	61.9	63.4
mus	51.8	100.0	100.0	96.4	100.0	100.0	99.8
mux	49.9	78.6	99.8	61.9	99.9	100.0	100.0
pmi	65.1	70.3	73.9	75.4	73.1	72.6	76.0
prt	24.9	34.5	42.5	50.8	41.6	39.8	43.7
seg	14.3	97.4	96.1	80.1	97.2	96.8	96.1
sick	93.8	96.1	96.3	93.3	98.4	97.0	96.7
soyb	13.5	89.5	90.3	92.8	91.4	90.3	76.2
tao	49.8	96.1	96.0	80.8	95.1	93.6	88.4
thy	19.5	68.1	65.1	80.6	92.1	92.1	86.3
veh	25.1	69.4	69.7	46.2	73.6	72.6	72.2
vote	61.4	92.4	92.6	90.1	96.3	96.5	95.4
vow	9.1	99.1	96.6	65.3	80.7	78.3	87.6
wne	39.8	95.6	96.8	97.8	94.6	92.9	96.3
zoo	41.7	94.6	92.5	95.4	91.6	92.5	92.6
Avg	44.8	80.0	82.4	78.0	82.1	81.8	81.7

Motivation

Alg. 4 is the winner in 8 problems with average 78.0

Alg. 2 is the winner for 4 problems with average 80.0

What is the best between both?



	Alg. 1	Alg. 2	Alg. 3	Alg. 4	Alg. 5	Alg. 6	Alg. 7
aud	25.3	76.0	68.4	69.6	79.0	81.2	57.7
aus	55.5	81.9	85.4	77.5	85.2	83.3	85.7
bal	45.0	76.2	87.2	90.4	78.5	81.9	79.8
bpa	58.0	63.5	60.6	54.3	65.8	65.8	68.2
bps	51.6	83.2	82.8	78.6	80.1	79.0	83.3
bre	65.5	96.0	96.7	96.0	95.4	95.3	96.0
cmc	42.7	44.4	46.8	50.6	52.1	49.8	52.3
gls	34.6	66.3	66.4	47.6	65.8	69.0	72.6
h-c	54.5	77.4	83.2	83.6	73.6	77.9	79.9
hep	79.3	79.9	80.8	83.2	78.9	80.0	83.2
irs	33.3	95.3	95.3	94.7	95.3	95.3	94.7
krk	52.2	89.4	94.9	87.0	98.3	98.4	98.6
lab	65.4	81.1	92.1	95.2	73.3	73.9	75.4
led	10.5	62.4	75.0	74.9	74.9	75.1	74.8
lym	55.0	83.3	83.6	85.6	77.0	71.5	79.0
mmg	56.0	63.0	65.3	64.7	64.8	61.9	63.4
mus	51.8	100.0	100.0	96.4	100.0	100.0	99.8
mux	49.9	78.6	99.8	61.9	99.9	100.0	100.0
pmi	65.1	70.3	73.9	75.4	73.1	72.6	76.0
prt	24.9	34.5	42.5	50.8	41.6	39.8	43.7
seg	14.3	97.4	96.1	80.1	97.2	96.8	96.1
sick	93.8	96.1	96.3	93.3	98.4	97.0	96.7
soyb	13.5	89.5	90.3	92.8	91.4	90.3	76.2
tao	49.8	96.1	96.0	80.8	95.1	93.6	88.4
thy	19.5	68.1	65.1	80.6	92.1	92.1	86.3
veh	25.1	69.4	69.7	46.2	73.6	72.6	72.2
vote	61.4	92.4	92.6	90.1	96.3	96.5	95.4
vow	9.1	99.1	96.6	65.3	80.7	78.3	87.6
wne	39.8	95.6	96.8	97.8	94.6	92.9	96.3
zoo	41.7	94.6	92.5	95.4	91.6	92.5	92.6
Avg	44.8	80.0	82.4	78.0	82.1	81.8	81.7

Motivation

We must use statistical tests for comparing the algorithms.

The problem:

How must I do the statistical experimental study?

What tests must I use?

	Alg. 1	Alg. 2	Alg. 3	Alg. 4	Alg. 5	Alg. 6	Alg. 7
aud	25.3	76.0	68.4	69.6	79.0	81.2	57.7
aus	55.5	81.9	85.4	77.5	85.2	83.3	85.7
bal	45.0	76.2	87.2	90.4	78.5	81.9	79.8
bpa	58.0	63.5	60.6	54.3	65.8	65.8	68.2
bps	51.6	83.2	82.8	78.6	80.1	79.0	83.3
bre	65.5	96.0	96.7	96.0	95.4	95.3	96.0
cmc	42.7	44.4	46.8	50.6	52.1	49.8	52.3
gls	34.6	66.3	66.4	47.6	65.8	69.0	72.6
h-c	54.5	77.4	83.2	83.6	73.6	77.9	79.9
hep	79.3	79.9	80.8	83.2	78.9	80.0	83.2
irs	33.3	95.3	95.3	94.7	95.3	95.3	94.7
krk	52.2	89.4	94.9	87.0	98.3	98.4	98.6
lab	65.4	81.1	92.1	95.2	73.3	73.9	75.4
led	10.5	62.4	75.0	74.9	74.9	75.1	74.8
lym	55.0	83.3	83.6	85.6	77.0	71.5	79.0
mmg	56.0	63.0	65.3	64.7	64.8	61.9	63.4
mus	51.8	100.0	100.0	96.4	100.0	100.0	99.8
mux	49.9	78.6	99.8	61.9	99.9	100.0	100.0
pmi	65.1	70.3	73.9	75.4	73.1	72.6	76.0
prt	24.9	34.5	42.5	50.8	41.6	39.8	43.7
seg	14.3	97.4	96.1	80.1	97.2	96.8	96.1
sick	93.8	96.1	96.3	93.3	98.4	97.0	96.7
soyb	13.5	89.5	90.3	92.8	91.4	90.3	76.2
tao	49.8	96.1	96.0	80.8	95.1	93.6	88.4
thy	19.5	68.1	65.1	80.6	92.1	92.1	86.3
veh	25.1	69.4	69.7	46.2	73.6	72.6	72.2
vote	61.4	92.4	92.6	90.1	96.3	96.5	95.4
vow	9.1	99.1	96.6	65.3	80.7	78.3	87.6
wne	39.8	95.6	96.8	97.8	94.6	92.9	96.3
zoo	41.7	94.6	92.5	95.4	91.6	92.5	92.6
Avg	44.8	80.0	82.4	78.0	82.1	81.8	81.7

Objective

To show some results on the use of statistical tests (nonparametric tests) for comparing algorithms in the fields of Data Mining and Computational Intelligence.

We will not discuss the performance measures that can be used neither the choice on the set of benchmarks.

Some guidelines on the use of appropriate nonparametrics tests depending on the situation will be given

OUTLINE (I)

- Introduction to Inferential Statistics
- Conditions for the safe use of parametric tests
 - Theoretical background
 - Checking the conditions in Data Mining Experiments
 - Checking the conditions in Parameter Optimization Experiments
- Basic non-parametric tests and case studies:
 - **■** For Pairwise Comparisons
 - For Multiple Comparisons involving control method
 - Data Mining: Neural Networks and Genetic Learning
 - **Evolutionary Algorithms: CEC'05 Special Session on Parameter Optimization**

OUTLINE (II)

- Advanced non-parametric tests and case studies:
 - For Multiple Comparisons involving control method
 - Post-hoc Procedures
 - Adjusted p-values
 - Detecting all pairwise differences in a multiple comparison
- Lessons Learned
 - Considerations on the use of nonparametric tests
 - **Recommendations on the use of nonparametric tests**
 - Frequent Questions
- Books of Interest and References
- Software

Website

http://sci2s.ugr.es/sicidm/





The web is organized according to the following **summary**:

- 1. Introduction to Inferential Statistics
- 2. Conditions for the safe use of Nonparametric Tests
- 3. Nonparametric tests
 - 3.1. Pairwise Comparisons
 - 3.2. Multiple Comparisons with a control method
 - 3.3. Multiple Comparisons among all methods
- 4. Case Studies
 - 4.1. Multiple Comparisons with a control method
 - 4.2. Multiple Comparisons among all methods
- 5. Considerations on the use of Nonparametric tests
- 6. Relevant Journal Papers with Data Mining and Computational Intelligence Case Studies
- Relevant books on Non-parametric tests
- 8. Topic Slides
- 9. Software and User's Guide

OUTLINE (I)

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Inferential Statistics

provide measures of how well your data (results of experiments) support your hypothesis and if your data are generalizable beyond what was tested (significance tests)

For example: Comparing two or various sets of experiments in a computational problem.

Parametric versus Nonparametric Statistics – When to use them and which is more powerful?

What is an hypothesis?

a prediction about a single population or about the relationship between two or more populations.

Hypothesis testing is procedure in which sample data are employed to evaluate a hypothesis.

The null hypothesis is a statement of no effect or no difference and it is expected to be rejected by the experimenter.

Examples of Null-Hypothesis

H_o: The 2 samples come from populations with the same distributions.

Or,

median of population 1 = median of population 2

(generalization with n samples)

Significance level a

- It is a confidence threshold that informs us whether or not to reject the null hypothesis.
- It must be pre-defined by the experimenter and a significance level of 90% (0.1) or 95% (0.05) is usually used, also 99% (0.01).

Significance level a

- If you decide for a significance level of 0.05 (95% certainty that there indeed is a significant difference), then a p-value (datum provided by the test) smaller than 0.05 indicates that you can reject the null-hypothesis.
- Remember: the null-hypothesis generally is associated to an hypothesis of equality or equivalence (equal means or distributions).
- So, if a test obtains p = 0.07, it means that you cannot reject the null hypothesis of equality ⇒ there is no significant differences in the analysis conducted

p-value

- Instead of stipulating a priori level of significance a, one could calculate the smallest level of significance that results in the rejection of the null hypothesis.
- This is the p-value, it provides information about "how significant" the result is.

It does it without committing to a particular level of significance.

There is at least one nonparametric test equivalent to a basic parametric test

Compare two variables

• If more than two variables

Parametric	Nonparametric
t-test	Sign test
	Wilcoxon signed rank test
ANOVA and	Friedman test
derivatives	and more
Tukey,	Bonferroni-Dunn,
Tamhane,	Holm, etc

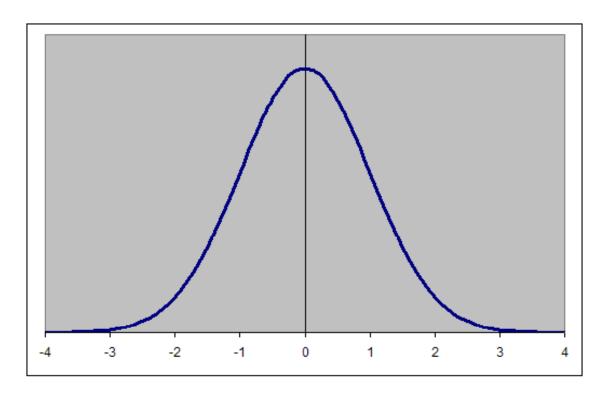
Parametric Assumptions

(t-test, ANOVA, ...)

- The observations must be independent
- Normality: The observations must be drawn from normally distributed populations
- Homoscedasticity: These populations must have the same variances

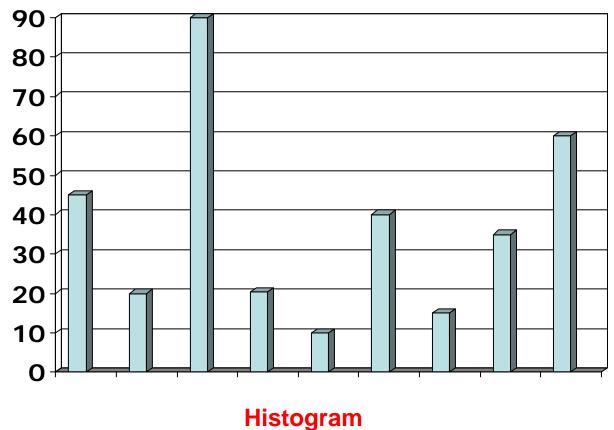
Normality Tip

If a histogram representing your data looks like this, you can conduct a parametric test!



Otherwise, don't conduct a parametric test!

The conclusions could be erroneous



Nonparametric Assumptions

(t-test, ANOVA, ...)

- The observations must be independent
- The data must be represented by ordinal numbering.

How do nonparametric tests work?

- ☐ Most nonparametric tests use *ranks* instead of raw data for their hypothesis testing.
- ☐ They apply a transformation procedure in order to obtain ranking data.

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Conditions for the safe use of parametric tests

- Theoretical background
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The distinction between parametric and nonparametric test is based on the level of measure represented by the data which will be analyzed.

A parametric test is able to use data composed by real values: But when we dispose of this type of data, we should not always use a parametric test.

There are some assumptions for a safe usage of parametric tests ad the non fulfillment of these conditions might cause a statistical analysis to lose credibility.

In order to use the parametric tests, is necessary to check the following conditions:

Independence: In statistics, two events are independent when the fact that one occurs does not modify the probability of the other one occurring.

- When we compare two optimization algorithms they are usually independent.
- When we compare two machine learning methods, it depends on the partition:
 - The independency is not truly verified in 10-fcv (a portion of samples is used either for training and testing in different partitions.
 - Hold out partitions can be safely take as independent, since training and test partitions do not overlap.

Parametric tests assume that the data are taken from normal distributions

Normality: An observation is normal when its behaviour follows a normal or Gauss distribution with a certain value of average μ and variance σ . A normality test applied over a sample can indicate the presence or absence of this condition in observed data.

- Kolmogorov-Smirnov
- Shapiro-Wilk
- D'Agostino-Pearson

- **Kolmogorov-Smirnov:** It compares the accumulated distribution of observed data with the accumulated Gaussian distribution expected.
- **Shapiro-Wilk:** It analyzes the observed data to compute the level of symmetry and kurtosis (shape of the curve) in order to compute the difference with respect to a Gaussian distribution afterwards.
- **D'Agostino-Pearson:** It computes the skewness and kurtosis to quantify how far from the Gaussian distribution is in terms of asymmetry and shape.

Heteroscedasticity: This property indicates the existence of a violation of the hypothesis of equality of variances.

Levene's test is used for checking if k samples present or not this homogeneity of variances (homoscedasticity).

Conditions for the safe use of parametric tests

- Theoretical background
- Checking the conditions in Data Mining Experiments
- Checking the conditions in Parameter Optimization Experiments

Checking the Conditions in Data Mining Experiments

FIRST CASE STUDY: Neural networks models:

MLP, RBFN (3 versions), LQV

Hold-Out Validation (HOV), 10FCV and 5x2CV (5 runs each one)

Data set	# Instances	# Attributes	# Classes
Breast	682	10	2
Cleveland	303	13	5
Crx	689	16	2
Glass	214	9	7
Iris	150	4	3
Pima	768	8	2
Wine	178	13	3
Wisconsin	699	10	2
Bupa	345	7	2
Lymphography	148	18	4
Monks	432	6	2
Page-blocks	5476	10	5
Pen-based	10992	16	10
Ringnorm	7400	20	2
Satimage	6435	36	7
Splice	3190	60	3

Checking the Conditions in Data Mining Experiments

TABLE 1. Kolmogorov-Smirnov test

Test of normality of Kolmogorov-Smirnov for HOV.

	Breast	Cleveland	Crx	Glass	Iris	Pima	Wine	Wisconsin
MLP	(00.)*	(.01)	(.20)	(.10)	(.00)	(00.)*	(.00)	(.00)
RBFN	(00.)	(.18)	(.07)	(.00)	(.00)	(.20)	(.01)	(.00)
RBFN Decremental	(00.)*	(.20)	(.00)	(.20)	(.00)	(.16)	(.00)	(.00)
RBFN Inc.	(.04)	(.20)	(.20)	(.01)	(.00)	(.03)	(.20)	(.00)
LVQ	(.11)	(.20)	(.04)	(.00)	(.04)	(.01)	(.07)	(.00)

Test of normality of Kolmogorov-Smirnov for 10FCV.

	Breast	Cleveland	Crx	Glass	Iris	Pima	Wine	Wisconsin
MLP	(.20)	(.17)	*(.00)	(.03)	(.00)	*(.01)	(.00)	(.00)
RBFN	(.02)	(.01)	(.20)	(.20)	(.00)	(.20)	(.00)	(.00)
RBFN Decremental	(.20)	(.20)	(.00)	(.20)	(.00)	(.18)	(.00)	(.00)
RBFN Inc.	(.10)	(.20)	(.20)	(.20)	(.00)	(.06)	(.03)	(.00)
LVQ	(.20)	(80.)	(.20)	(.20)	(.20)	(.20)	(.00)	(.00)

Test of Normality of Kolmogorov-Smirnov for 5 x 2CV.

	Breast	Cleveland	Crx	Glass	Iris	Pima	Wine	Wisconsin
MLP	(.18)	(.20)	(.20)	*(.04)	(.20)	(.20)	*(.04)	(.20)
RBFN	(.20)	(.20)	(.09)	(.00)	(.20)	(.20)	(.00)	(.01)
RBFN Decremental	(.00)	(.05)	(.00)	(.00)	(.00)	(.20)	(.00)	(.01)
RBFN Inc.	(.01)	(.20)	(.20)	(.20)	(.01)	(.20)	(.20)	(.04)
LVQ	(.20)	(.04)	(.05)	(.07)	(.03)	(.05)	(.00)	(.07)

Checking the Conditions in Data Mining Experiments

TABLE 2. Comparison among validations

Test of Normality of D'Agostino-Pearson for HOV.

	Breast	Cleveland	Crx	Glass	Iris	Pima	Wine	Wisconsin
MLP	(.17)	(.65)	(.06)	(.14)	(.00)	(.35)	(.01)	(.12)
RBFN	(.00)	(.18)	(.38)	(.02)	(.00)	(.59)	(.01)	(.26)
RBFN Decremental	(.00)	(.88)	(.00)	(.10)	(.00)	(.43)	(.40)	(.00)
RBFN Inc.	(.24)	(.06)	(.50)	(.09)	(.09)	(.10)	(.94)	(.98)
LVQ	(.31)	(.59)	(.11)	(00.)*	(.21)	(.00)	(.05)	(.00)

Test of Normality of D'Agostino-Pearson for 10FCV.

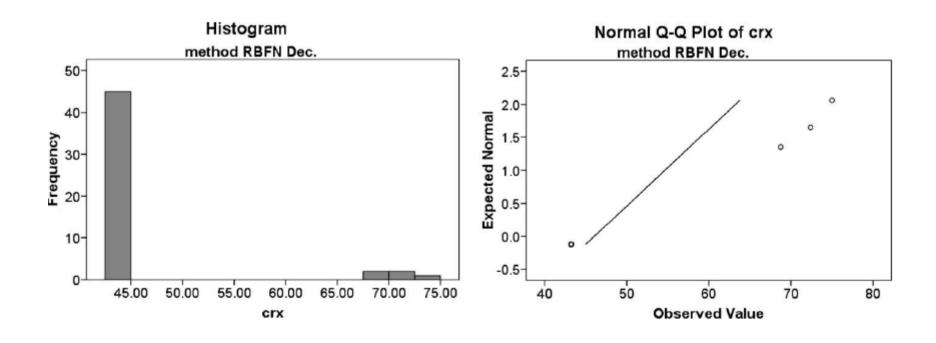
	Breast	Cleveland	Crx	Glass	Iris	Pima	Wine	Wisconsin
MLP	(.21)	(.70)	(.00)	(.51)	(.03)	(.06)	(.03)	(.00)
RBFN	(.63)	(.20)	(.61)	(.60)	(.00)	(.27)	(00.)	(.03)
RBFN Decremental	(.06)	(.56)	(.00)	(.63)	(.15)	(.10)	(.00)	(.39)
RBFN Inc.	(.36)	(.65)	(.90)	(.11)	(.38)	(.04)	(.53)	(.07)
LVQ	(.78)	(.00)	(.02)	(1.00)	(.18)	(.23)	(00.)*	(.00)

Test of Normality of D'Agostino-Pearson for 5 x 2CV.

	Breast	Cleveland	Crx	Glass	Iris	Pima	Wine	Wisconsin
MLP	(.92)	(.60)	*(.03)	(.53)	(.11)	(.46)	(.53)	(.14)
RBFN	(.90)	(.63)	*(.22)	(.02)	(.03)	(.06)	(.11)	(.02)
RBFN Decremental	(.00)	(.17)	(.00)	(.11)	(.00)	(.82)	(.02)	(.25)
RBFN Inc.	(.02)	(.34)	(.34)	(.90)	(.56)	(.18)	(.90)	(.66)
LVQ	(.42)	(.09)	(.11)	(.65)	(.30)	(.76)	(.03)	(00.)*

Checking the Conditions in Data Mining Experiments

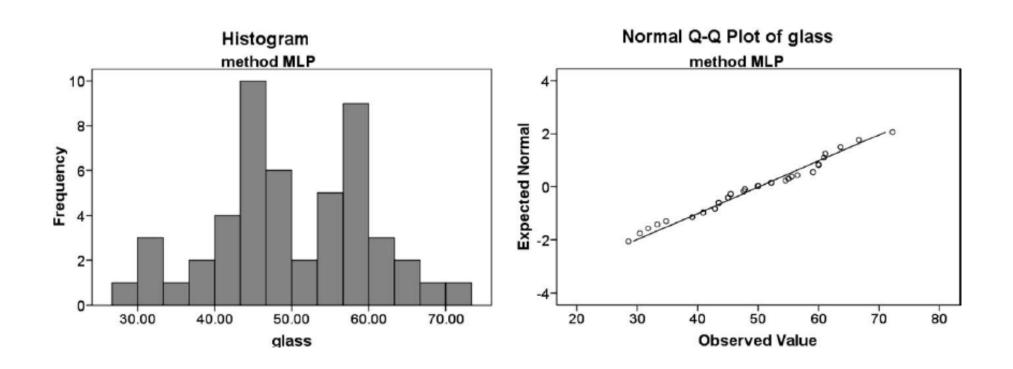
Histograms and Q-Q Grapics



* A Q-Q graphic represents a confrontation between the quartiles from data observed and those from the normal distributions. Absolute lack of normality. 33

Checking the Conditions in Data Mining Experiments

Histograms and Q-Q Grapics



Checking the Conditions in Data Mining Experiments

TABLE 3. Test of HETEROSCEDASTICITY OF LEVENE (BASED ON MEANS)

	Breast	Cleveland	Crx	Glass	Iris	Pima	Wine	Wisconsin
HOV	(.00)	(.00)	(.00)	(.00)	(.00)	(.00)	(.00)	(00.)*
10FCV	(.00)	(.00)	(.00)	(.00)	(.00)	(.20)	(.00)	(.01)
5 × 2CV	(.00)	(.01)	(00.)	(.00)	(.00)	(.00)	(00.)	(00.)

Table 3 shows the results by applying Levene's tests, where the symbol "*" indicates that the variances of the distributions of the different algorithms for a certain function are not homogeneities (we reject the null hypothesis).

Checking the Conditions in Data Mining Experiments

SECOND CASE STUDY: Genetics-Based Machine Learning

- We have chosen four Genetic Interval Rule Based Algorithms:
 - Pittsburgh Genetic Interval Rule Learning Algorithm.
 - XCS Algorithm.
 - GASSIST Algorithm.
 - HIDER Algorithm.
- GBML will be analyzed by two performance measures: Accuracy and Cohen's kappa.
- How we state which is the best?

Checking the Conditions in Data Mining Experiments

Experimental Study

We have selected 14 data sets from UCI repository.

Data set	#Ex.	#Atts.	#C.
bupa (bup)	345	6	2
cleveland (cle)	297	13	5
ecoli (eco)	336	7	8
glass (gla)	214	9	7
haberman (hab)	306	3	2
iris (iri)	150	4	3
monk-2 (mon)	432	6	2
new-Thyroid (new)	215	5	3
pima (pim)	768	8	2
vehicle (veh)	846	18	4
vowel (vow)	988	13	11
wine (win)	178	13	3
wisconsin (wis)	683	9	2
yeast (yea)	1484	8	10

Conditions for the Safe Use of Parametric Tests Checking the Conditions in Data Mining Experiments

TABLE I. Normality condition in accuracy

Shapiro-Wilk															
	bup cle eco gla hab iri mon new pim veh vow win wis yea Pitts-GIRLA * (.02) * (.00														
Pitts-GIRLA * (.02) * (.00) * (.00) (.73) * (.00) * (.00) * (.00) * (.01) * (.00) * (.00)												* (.00)	* (.00)	* (.00)	
XCS (.25) * (.03) (.23) * (.00) * (.02) * (.00) * (.00) * (.00) * (.00) * (.03) (.17) (.30) * (.00) * (.00)												(.45)			
GASSIST (.39) (.21) (.07) (.19) * (.04) * (.00) (.07) * (.00) (.12) (.81) (.51) * (.00) * (.00) (.83)													(.83)		
HIDER	(.11)	(.42)	(.22)	* (.00)	* (.01)	* (.00)	(.06)	* (.00)	* (.00)	(.25)	(.15)	* (.00)	* (.00)	(.23)	
						D'Agost	ino-Pearso	n							
	bup	cle	eco	gla	hab	iri	mon	new	pim	veh	vow	win	wis	yea	
Pitts-GIRLA	(.13)	(.10)	* (.00)	(.69)	* (.00)	(.11)	* (.00)	(.71)	* (.00)	* (.02)	* (.00)	* (.00)	* (.00)	* (.00)	
XCS	(.44)	(.09)	(.61)	(.06)	(.22)	(.06)	* (.00)	* (.00)	(.24)	(.33)	(.40)	* (.00)	* (.03)	(.48)	
GASSIST	(.55)	(.75)	(.59)	(.42)	(.79)	(.19)	(.89)	(.89)	(.25)	(.65)	(.18)	* (.03)	* (.03)	(.95)	
HIDER	(.07)	(.52)	(.42)	(.05)	(.78)	* (.00)	(.19)	* (.00)	* (.00)	(.43)	(.37)	* (.00)	* (.02)	(.18)	

a value smaller than 0.05 indicates that you can reject the **null-hypothesis** (i.e. the normality condition is not satisfied) and it is noted with "*"

Conditions for the Safe Use of Parametric Tests Checking the Conditions in Data Mining Experiments

GBML Case of Study: some facts

- Conditions needed for the application of parametric tests are not fulfilled in some cases.
 - The size of the sample should be enough (50)
- One main factor: the nature of the problem
- Graphically, we can use Q-Q graphics and histograms to see the normality

Checking the Conditions in Data Mining Experiments

Analyzing parametric tests

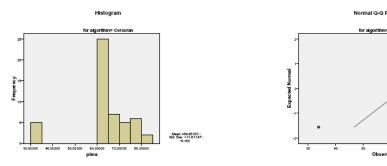
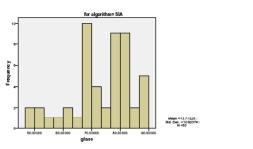


Figure 1: Results of Pitts-GIRLA over pima data set in 10fcv: Histogram and Q-Q Graphic.



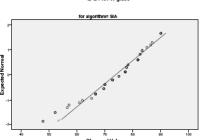


Figure 2: Results of SIA over glass data set in 10fcv: Histogram and Q-Q Graphic.

* A Q-Q graphic represents a confrontation between the quartiles from data observed and those from the normal distributions.

Checking the Conditions in Data Mining Experiments

TABLE 2. Test of HETEROSCEDASTICITY OF LEVENE

(BASED ON MEANS)

	bup	cle	eco	gla	hab	iri	mon	new	pim	veh	VOW	win	wis	yea
Accuracy	(.13)	* (.00)	(.36)	(.34)	* (.01)	(.40)	* (.00)	(.26)	(.16)	* (.00)	* (.03)	* (.00)	* (.00)	* (.00)
Cohen's kappa	(.51)	(.05)	(.39)	(.25)	* (.04)	(.40)	* (.00)	(.40)	* (.00)	* (.00)	* (.03)	* (.00)	* (.00)	* (.00)

Table 2 shows the results by applying Levene's tests, where the symbol "*" indicates that the variances of the distributions of the different algorithms for a certain function are not homogeneities (we reject the null hypothesis).

Checking the Conditions in Data Mining Experiments

NN and GBML do not verify parametric conditions.

Similar studies can be presented with them.

J. Luengo, S. García, F. Herrera, A Study on the Use of Statistical Tests for Experimentation with Neural Networks: Analysis of Parametric Test Conditions and Non-Parametric Tests. Expert Systems with Applications 36 (2009) 7798-7808 doi:10.1016/j.eswa.2008.11.041.

S. García, A. Fernandez, J. Luengo, F. Herrera, A Study of Statistical Techniques and Performance Measures for Genetics-Based Machine Learning: Accuracy and Interpretability. Soft Computing 13:10 (2009) 959-977, doi:10.1007/s00500-008-0392-y.

Statistical Analysis of Experiments in Data Mining and Computational Intelligence

Conditions for the safe use of parametric tests

- Theoretical background
- Checking the conditions in Data Mining Experiments
- Checking the conditions in Parameter Optimization Experiments

Checking the Conditions in Parameter Optimization Experiments

Special Session on Real-Parameter Optimization at CEC-05, Edinburgh, UK, 2-5 Sept. 2005

25 functions with real parameters, 10 variables: f1-f5 unimodal functions f6-f25 multimodal functions

P. N. Suganthan, N. Hansen, J. J. Liang, K. Deb, Y. P. Chen, A. Auger, and S. Tiwari, "Problem definitions and evaluation criteria for the CEC 2005 special session on real parameter optimization." Nanyang Technological University, Tech. Rep., 2005, available as http://www.ntu.edu.sg/home/epnsugan/index_files/CEC-05/Tech-Report-May-30-05.pdf.

N. Hansen, "Compilation of Results on the CEC Benchmark Function Set," Institute of Computational Science, ETH Zurich, Switerland, Tech. Rep., 2005, available as http://www.ntu.edu.sg/home/epnsugan/index_files/CEC-05/compareresults.pdf.

Source: S. García, D. Molina, M. Lozano, F. Herrera, A Study on the Use of Non-Parametric Tests for Analyzing the Evolutionary Algorithms' Behaviour: A Case Study on the CEC'2005 Special Session on Real Parameter Optimization. *Journal of Heuristics* 15 (2009) 617-644, doi: 10.1007/s10732-008-9080-4.

Checking the Conditions in Parameter Optimization Experiments

- Algorithms involved in the comparison:
 - BLX-GL50 (Garcia-Martinez & Lozano, 2005): Hybrid Real-Coded Genetic Algorithms with Female and Male Differentiation
 - BLX-MA (Molina et al., 2005): Adaptive Local Search Parameters for Real-Coded Memetic Algorithms
 - CoEVO (Posik, 2005): Mutation Step Co-evolution
 - **DE** (**Ronkkonen** *et al.*,2005):Differential Evolution
 - DMS-L-PSO: Dynamic Multi-Swarm Particle Swarm Optimizer with Local Search
 - EDA (Yuan & Gallagher, 2005): Estimation of Distribution Algorithm
 - G-CMA-ES (Auger & Hansen, 2005): A restart Covariance Matrix Adaptation Evolution Strategy with increasing population size
 - K-PCX (Sinha et al., 2005): A Population-based, Steady-State real-parameter optimization algorithm with parent-centric recombination operator, a polynomial mutation operator and a niched -selection operation.
 - L-CMA-ES (Auger & Hansen, 2005): A restart local search Covariance Matrix Adaptation Evolution Strategy
 - L-SaDE (Qin & Suganthan, 2005): Self-adaptive Differential Evolution algorithm with Local Search
 - SPC-PNX (Ballester et al.,2005): A steady-state real-parameter GA with PNX crossover operator

Conditions for the Safe Use of Parametric Tests Checking the Conditions in Parameter Optimization Experiments

Table 1 Test of normality of Kolmogorov-Smirnov

	f1	f2	f3	f4	f5	f6	f7	f8	f9
BLX-GL50 BLX-MA	(.20) * (.01) f10	* (.04) * (.00) f11	* (.00) * (.01) f12	(.14) * (.00) f13	* (.00) * (.00) f14	* (.00) (.16)	* (.04) (.20)	(.20) * (.00) f17	* (.00) * (.00) f18
BLX-GL50 BLX-MA	(.10)	(.20)	* (.00) * (.00)	(.20)	(.20)	* (.00) * (.00)	* (.00)	(.20)	* (.00) * (.00)
	f19	f20	f21	f22	f23	f24	f25		
BLX-GL50 BLX-MA	* (.00) * (.00)	* (.00) * (.00)	* (.00) * (.00)	* (.00) * (.00)	* (.00) * (.00)	* (.00) * (.00)	* (.00) * (.02)		

Conditions for the Safe Use of Parametric Tests Checking the Conditions in Parameter Optimization Experiments

Table 3 Test of normality of D'Agostino-Pearson

	f1	f2	f3	f4	f5	f6	f7	f8	f9
BLX-GL50	(.10)	(.06)	* (.00)	(.24)	* (.00)	* (.00)	(.28)	(.21)	* (.00)
BLX-MA	* (.00)	* (.00)	(.22)	* (.00)	* (.00)	* (.00)	(.19)	(.12)	* (.00)
	f10	f11	f12	f13	f14	f15	f16	f17	f18
BLX-GL50	(.17)	(.19)	* (.00)	(.79)	(.47)	* (.00)	* (.00)	(.07)	* (.03)
BLX-MA	(.89)	* (.00)	* (.03)	(.38)	(.16)	* (.00)	(.21)	(.54)	* (.04)
	f19	f20	f21	f22	f23	f24	f25		
BLX-GL50	(.05)	(.05)	(.06)	* (.01)	* (.00)	* (.00)	(.11)		
BLX-MA	* (.00)	* (.00)	(.25)	* (.00)	* (.00)	* (.00)	(.20)		

Checking the Conditions in Parameter Optimization Experiments

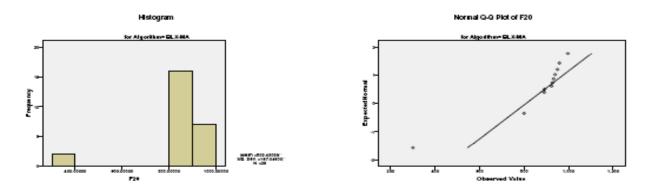


Figure 1: Example of non-normal distribution: Function f20 and BLX-GL50 algorithm: Histogram and Q-Q Graphic.

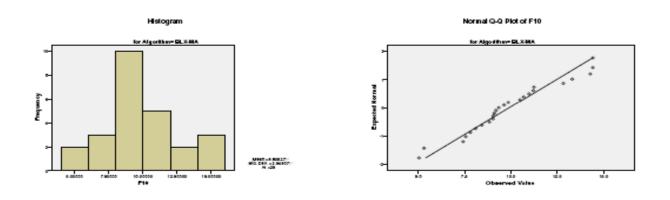


Figure 2: Example of normal distribution: Function f10 and BLX-MA algorithm: Histogram and Q-Q Graphic.

Conditions for the Safe Use of Parametric Tests Checking the Conditions in Parameter Optimization Experiments

Table 4 Test of heteroscedasticity of Levene (based on means)

	f1	f2	f3	f4	f5	f6	f7	f8	f9
LEVENE	(.07)	(.07)	* (.00)	* (.04)	* (.00)	* (.00)	* (.00)	(.41)	* (.00)
	f10	f11	f12	f13	f14	f15	f16	f17	f18
LEVENE	(.99)	* (.00)	(.98)	(.18)	(.87)	* (.00)	* (.00)	(.24)	(.21)
LEVENE	(.99) f19			(.18) f22			* (.00) f25	(.24)	(.21)

Statistical Analysis of Experiments in Data Mining and Computational Intelligence

OUTLINE (I)

- Introduction to Inferential Statistics
- Conditions for the safe use of parametric tests
 - Theoretical background
 - Checking the conditions in Data Mining Experiments
 - **■** Checking the conditions in Parameter Optimization Experiments
- Basic non-parametric tests and case studies:
 - **■** For Pairwise Comparisons
 - For Multiple Comparisons involving control method
 - Data Mining: Neural Networks and Genetic Learning
 - **Evolutionary Algorithms: CEC'05 Special Session on Parameter Optimization**

Statistical Analysis of Experiments in Data Mining and Computational Intelligence

Basic Non-Parametric Tests and Case Studies

- For Pairwise Comparisons
- For Multiple Comparisons involving a Control Method
- Data Mining: Neural Networks and Genetic Learning
- **Evolutionary Algorithms: CEC'05 Special Session of Parameter Optimization**

Pairwise Comparisons involve Two-Sample Tests

When comparing means of two samples to make inferences about differences between two populations, there are 4 main tests that could be used:

	Unpaired data	Paired data
Parametric test	Independent-Samples / T-Test	Paired-Samples T-Test
Non-parametric test	Mann-Whitney U test (or Wilcoxon rank-sum test)	Wilcoxon Signed-Ranks test (Also, Sign test)

Count of Wins, Losses and Ties: The Sign Test

It a classic form of inferential statistics that use the binomial distribution. If two algorithms compared are, assumed under the null-hypothesis, equivalent, each should win approximately N/2 out of N datasets/problems.

The number of wins are distributed following a binomial distribution. The critical number of wins are presented in the following Table for α =0.05 and α =0.1:

#data sets	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25
W _{0.05}	5	6	7	7	8	9	9	10	10	11	12	12	13	13	14	15	15	16	17	18	18
$W_{0.10}$	5	6	6	7	7	8	9	9	10	10	11	12	12	13	13	14	14	15	16	16	17

The Sign Test

For a greater number of datasets/problems, the number of wins is under the null-hypothesis distributed according to $N(N/2, \sqrt{N}/2)$.

Thus, if an algorithm obtains a number of wins which is at least

 $N/2 + 1.96\sqrt{N}/2$ the algorithm is significantly better with $\alpha = 0.05$. Tieds are split between the two algorithms. If they are an odd number, one is ignored.

Example of the Sign Test

dataset	C4.5	C4.5m	Sign	
Adult	0.763	0.768	+	Classification problem with 14 datasets.
Breast	0.599	0.591	-	C4.5 standard vs C4.5 with m parameter
Wisconsin	0.954	0.971	+	(minimum number of examples for creating
Cmc	0.628	0.661	+	a leaf) tuned for AUC measure.
Ionosphere	0.882	0.888	+	
Iris	0.936	0.931	-	Number of wins of $C4.5m = 10$
Bupa	0.661	0.668	+	
Lung	0.583	0.583	=	Number of loses of $C4.5m = 2$
Lymphography	0.775	0.838	+	Number of ties $= 2$
Mushroom	1.000	1.000	=	
Tumor	0.940	0.962	+	
Rheum	0.619	0.666	+	Moreover, one tie is added in the wins
Voting	0.972	0.981	+	count. No. of wins $= 11$.
Wine	0.957	0.978	+	

Example of Sign Test

According to the previous Table, this difference is significant with $\alpha = 0.05$.

#data sets																					
W _{0.05}	5	6	7	7	8	9	9	10	10	11	12	12	13	13	14	15	15	16	17	18	18
$W_{0.10}$	5	6	6	7	7	8	9	9	10	10	11	12	12	13	13	14	14	15	16	16	17

This test does not assume any commensurability of scores or differences nor does it assume normal distributions and is thus applicable to any data. On the other hand, it is much weaker than the Wilcoxon signed-ranks test because it will not reject the null-hypothesis unless one algorithm almost always outperforms the other.

Wilcoxon Signed-Ranks Test for Paired Samples

The Wilcoxon Signed-Ranks test is used in exactly the same situations as the paired t-Test (i.e., where data from two samples are paired).

In general, the Test asks:

 H_0 : The 2 samples come from populations with the same distributions. Or, median of population 1 = median of population 2

The test statistic is based on ranks of the differences between pairs of data.

NOTE: If you have \leq 5 pairs of data points, the Wilcoxon Signed-Ranks test can never report a 2-tailed p-value < 0.05

Procedure for the Wilcoxon Signed-Ranks Test

- 1. For each pair of data, calculate the difference. Keep track of the sign (+ve or -ve).
- 2. Temporarily ignoring the sign of the difference, rank the absolute values of the difference. When the differences have the same value, assign them the mean of the ranks involved in the tie.
- 3. Consider the sign of the differences again and ADD up the ranks of all the positive differences and all the negative differences
- (R⁺, R⁻). Ranks of difference equal to 0 are split evenly among the sums; if there is an odd number of them, one is ignored.

Procedure for the Wilcoxon Signed-Ranks Test

4. Let T be the <u>smaller</u> of the sums of positive and negative differences. $T = Min \{R^+, R^-\}$.

Use an appropriate Statistical Table or computer to determine the test statistic, critical region or P-values.

- 5. Reject the H_0 if test statistic \leq critical value, or if $P \leq \alpha$ (alpha).
- 6. Report Test results.

Example of the Wilcoxon Signed-Ranks Test

	_				
dataset	C4.5	C4.5m	Difference	Rank	_
Adult	0.763	0.768	+0.005	3.5	_
Breast	0.599	0.591	-0.008	7	
Wisconsin	0.954	0.971	+0.017	9	$R^+ = 3.5 + 9 + 12 + 5 + 6 + 14 +$
Cmc	0.628	0.661	+0.033	12	11 + 13 + 8 + 10 + 1.5 = 93
Ionosphere	0.882	0.888	+0.006	5	
Iris	0.936	0.931	-0.005	3.5	
Bupa	0.661	0.668	+0.007	6	
Lung	0.583	0.583	0.000	1.5	
Lymphograph	0.775	0.838	+0.063	14	$R^{-} = 7 + 3.5 + 1.5 = 12$
Mushroom	1.000	1.000	0.000	1.5	
Tumor	0.940	0.962	+0.022	11	
Rheum	0.619	0.666	+0.047	13	
Voting	0.972	0.981	+0.009	8	60
Wine	0.957	0.978	+0.021	10	

Example of the Wilcoxon Signed-Ranks Test

\mathbf{P}^+	= 3	5 +	Q +	- 12	+ 5	+
$\boldsymbol{\Gamma}$	$ \mathcal{I}$.ງ⊤	ファ	T 1 🚄	\top)	

$$6+14+11+13+$$

$$8 + 10 + 1.5 = 93$$

$$R^{-} = 7 + 3.5 + 1.5 = 12$$

$$T = Min \{R^+, R^-\} = 12$$

$$\alpha$$
= 0.05, N = 14 dif = 21

	LEVEL OF SIGNIFICANCE FOR ONE-TAIL			ED TEST
	n	0.025	0.01	0.005
		LEVEL OF SIGNIFIC	CANCE FOR TWO-TAI	LED TEST
		0.05	0.02	0.01
	6	0	_	_
	7	2	0	_
	8	4	2	0
	9	6	3	2
	10	8	5	3
	11	11	7	5
	12	14	10	7
	13	17	13	10
	14	21	16	13
	15	25	20	16
	16	30	24	20
	17	35	28	23
	18	40	33	28
	19	46	38	32
	20	52	43	38
	21	59	49	43
	22	66	56	49
	23	73	62	55
	24	81	69	61
	25	89	77	68 61

Example of the Wilcoxon Signed-Ranks Test

Critical value for T for N up to 25.

It T <= dif (table-value) then Reject the H_o

	LEVEL OF SIGNIFICANCE FOR ONE-TAILED TEST			
n	0.025	0.01	0.005	
	LEVEL OF SIGNIFICANCE FOR TWO-TAILED TEST			
	0.05	0.02	0.01	
6	0	_	_	
7	2	0	_	
8	4	2	0	
9	6	3	2	
10	8	5	3	
11	11	7	5	
12	14	10	7	
13	17	13	10	
14	21	16	13	
15	25	20	16	
16	30	24	20	
17	35	28	23	
18	40	33	28	
19	46	38	32	
20	52	43	38	
21	59	49	43	
22	66	56	49	
23	73	62	55	
24	81	69	61	
25	89	77	68	

For $n \le 30$: use T values (and refer to a Table B.12. Critical Values of the Wilcoxon T Distribution, Zar, App 101)

For n > 30: use z-scores (z is distributed approximately normally). (and refer to the z-Table, Table B.2. Zar – Proportions of the Normal Curve (One-tailed), App 17)

$$z = \frac{T - \frac{n(n+1)}{4}}{\sqrt{\frac{n(n+1)(2n+1)}{24}}}$$

Where,

with $\alpha = 0.05$, the null-hypothesis can be rejected if z is smaller than -1.96.

The Wilcoxon signed ranks test is more sensible than the ttest. It assumes commensurability of differences, but only qualitatively: greater differences still count more, which is probably desired, but the absolute magnitudes are ignored.

From the statistical point of view, the test is safer since it does not assume normal distributions. Also, the outliers (exceptionally good/bad performances on a few data-sets/problems) have less effect on the Wilcoxon than on the t-test.

The Wilcoxon test assumes continuous differences, therefore they should not be rounded to one or two decimals, since this would decrease the power of the test due to a high number of ties.

Wilcoxon Signed-Ranks Test in SPSS

Analyze \rightarrow Nonparametric Tests \rightarrow 2 Related Samples Tests

• Select pair(s) of variables

Select Wilcoxon

Wilcoxon Signed-Ranks Test in SPSS

OUTPUT

		Ν	Mean Rank	Sum of Ranks
beta-endorphin	Negative Ranks	0 ^a	.00	.00
conc. after (pmol/l) -	Positive Ranks	11 ^b	6.00	66.00
beta-endorphin	Ties	0 ^c		
conc. before (pmol/l)	Total	11		

a. beta-endorphin conc. after (pmol/l) < beta-endorphin conc. before (pmol/l)

Test Statistics 1

	beta-endorphi	
n conc. after		
	(pmol/l) -	
	beta-endorphi	
	n conc. before	
	(pmol/l)	
Z	-2.934 ^a	
Asymp. Sig. (2-tailed)	.003	

a. Based on negative ranks.

b. beta-endorphin conc. after (pmol/l) > beta-endorphin conc. before (pmol/l)

C. heta-endorphin conc. hefore (pmol/l) = heta-endorphin conc. after (pmol/l)

b. Wilcoxon Signed Ranks Test

Statistical Analysis of Experiments in Data Mining and Computational Intelligence

Basic Non-Parametric Tests and Case Studies

- **■** For Pairwise Comparisons
- For Multiple Comparisons involving a Control Method
- Data Mining: Neural Networks and Genetic Learning
- Evolutionary Algorithms: CEC'05 Special Session of Parameter Optimization

Using Wilcoxon test for comparing multiple pairs of algorithms:

Wilcoxon's test performs individual comparisons between two algorithms (pairwise comparisons). The *p-value in a pairwise comparison is independent from another* one. If we try to extract a conclusion involving more than one pairwise comparison in a Wilcoxon's analysis, we will obtain an accumulated error coming from the combination of pairwise comparisons. In statistical terms, we are losing the control on the Family Wise Error Rate (FWER), defined as the probability of making one or more false discoveries among all the hypotheses when performing multiple pairwise tests.

Advanced Non-Parametric Tests and Case For Multiple Comparison Vivolving a Control Method

When a p-value is considered in a multiple comparison, it reflects the probability error of a certain comparison, but it does not take into account the remaining comparisons belonging to the family.

If one is comparing k algorithms and in each comparison the level of significance is α , then in a single comparison the probability of not making a Type I error is $(1-\alpha)$, then the probability of not making a Type I error in the k-1 comparison is $(1-\alpha)\cdot(k-1)$. Then the probability of making one or more Type I error is $1-(1-\alpha)\cdot(k-1)$.

For instance, if $\alpha = 0.05$ and k = 10, this is 0.37, which is rather high.

Using Wilcoxon test for comparing multiple pairs of algorithms:

The true statistical signification for the pairwise comparison test is given by:

$$p = P(Reject \ H_0|H_0 \ true) =$$

$$= 1 - P(Accept \ H_0|H_0 \ true) =$$

$$= 1 - P(Accept \ A_k = A_i, i = 1, ..., k - 1|H_0 \ true) =$$

$$= 1 - \prod_{i=1}^{k-1} P(Accept A_k = A_i|H_0 \ true) =$$

$$= 1 - \prod_{i=1}^{k-1} [1 - P(Reject \ A_k = A_i|H_0 \ true)] =$$

$$= 1 - \prod_{i=1}^{k-1} (1 - p_{H_i})$$

Using Multiple Comparison Procedures:

Making pairwise comparisons allows us to conduct this analysis, but the experiment wise error can not be previously controlled. Furthermore, a pairwise comparison is not influenced by any external factor, whereas in a multiple comparison, the set of algorithms chosen can determine the results of the analysis.

Multiple comparison procedures are designed for allowing us to fix the FWER before performing the analysis and for taking into account all the influences that can exist within the set of results for each algorithm.

Multiple Comparison Procedures:

Parametric	Nonparametric
ANOVA	Friedman's test
	Iman-Davenport's test
Turkey, Dunnet,	Bonferroni-Dunn's test Holm's method Hochberg's method

Friedman's test: It is a non-parametric equivalent of the test of repeated-measures ANOVA. It computes the ranking of the observed results for algorithm (r_j) for the algorithm j with k algorithms) for each function/algorithm, assigning to the best of them the ranking 1, and to the worst the ranking k.

Under the null hypothesis, formed from supposing that the results of the algorithms are equivalent and, therefore, their rankings are also similar, the Friedman statistic

$$\chi_F^2 = \frac{12N}{k(k+1)} \left[\sum_j R_j^2 - \frac{k(k+1)^2}{4} \right]$$

is distributed according to $\operatorname{con} k$ - 1 degrees of freedom, being, $R_j = \frac{1}{N} \sum_i r_i^j$ and N the number of functions/algorithms. (N > 10, k > 5) (Table B.1. Critical Values of the Chi-Square Distribution, App. 12, Zar).

Iman and Davenport's test: It is a metric derived from the Friedman's statistic given that this last metric produces a conservative undesirably effect. The statistic is:

$$F_F = \frac{(N-1)\chi_F^2}{N(k-1) - \chi_F^2}$$

and it is distributed according to a F distribution with k-1 and (k-1)(N-1) degrees of freedom.

(Table B.4. Critical values of the F Distribution, App. 21, Zar).

For Multiple Comparisons involving a Control Method

Example of the Friedman Test

The results obtained (performances) are arranged by a matrix of data with data sets in the rows and algorithms in the columns.

C4.5 with cf parameter is the version which optimizes AUC considering various levels of confidence for pruning a leaf.

dataset	C4.5	C4.5m	C4.5cf	C4.5cf,m
Adult	0.763	0.768	0.771	0.798
Breast	0.599	0.591	0.590	0.569
Wisconsin	0.954	0.971	0.968	0.967
Cmc	0.628	0.661	0.654	0.657
Ionosphere	0.882	0.888	0.886	0.898
Iris	0.936	0.931	0.916	0.931
Bupa	0.661	0.668	0.609	0.685
Lung	0.583	0.583	0.563	0.625
Lymphography	0.775	0.838	0.866	0.875
Mushroom	1.000	1.000	1.000	1.000
Tumor	0.940	0.962	0.965	0.962
Rheum	0.619	0.666	0.614	0.669
Voting	0.972	0.981	0.975	0.975
Wine	0.957	0.978	0.946	0.970
Lymphography Mushroom Tumor Rheum Voting	0.775 1.000 0.940 0.619 0.972	0.838 1.000 0.962 0.666 0.981	0.866 1.000 0.965 0.614 0.975	0.875 1.000 0.962 0.669 0.975

For Multiple Comparisons involving a Control Method

Example of the Friedman Test

Rankings are assigned in increasing order from the best to the worst algorithm for each dataset/problem.

Ties in performance are computed by averaged rankings.

The most interesting datum for now is the *Average Rank* for each algorithm.

dataset	C4.5	C4.5m	C4.5cf	C4.5cf,m
Adult	4	3	2	1
Breast	1	2	3	4
Wisconsin	4	1	2	3
Cmc	4	1	3	2
Ionosphere	4	2	3	1
Iris	1	2.5	4	2.5
Bupa	3	2	4	1
Lung	2.5	2.5	4	1
Lymphography	4	3	2	1
Mushroom	2.5	2.5	2.5	2.5
Tumor	4	2.5	1	2.5
Rheum	3	2	4	1
Voting	4	1	2	3
Wine	3	1	4	2
Average Rank	3.143	2.000	2.893	1.964 ₇₆

For Multiple Comparisons involving a Control Method

	C4.5	C4.5m	C4.5cf	C4.5cf,m
Average Rank	3.143	2.000	2.893	1.964

Friedman's measure

$$\chi_F^2 = \frac{12N}{k(k+1)} \left[\sum_j R_j^2 - \frac{k(k+1)^2}{4} \right] =$$

$$= \frac{12\cdot14}{4\cdot5} \left[9.878 + 4.000 + 8.369 + 3.857 - \frac{4\cdot25}{4} \right] =$$

$$= 9.28$$

Observing the critical value, it can be concluded that it rejects the null hypothesis

For Multiple Comparisons involving a Control Method

	C4.5	C4.5m	C4.5cf	C4.5cf,m
Average Rank	3.143	2.000	2.893	1.964

Iman and Davenport's measure

$$F_F = \frac{(N-1)\chi_F^2}{N(k-1) - \chi_F^2} = \frac{13.9.28}{13.3 - 9.28} = 3.69$$

$$F_F = 3.69, F(3,3x13) = 2.85$$

Observing the critical value, it can be concluded that it rejects the null hypothesis

If the null hypothesis is rejected by Friedman or Iman-Davenport test, we can proceed with a post-hoc test:

The most frequent case is when we want to compare one algorithm (the proposal) with a set of algorithm. This type of comparison involves a CONTROL method, and it is usually denoted as a 1 x n comparison.

The simplest procedure in $1 \times n$ comparisons is the Bonferroni-Dunn test. It adjusts the global level of significance by dividing it by (k-1) in all cases, being k the number og algorithms.

The performance of two algorithms is significantly different if the corresponding average ranks differ by at least the critical difference:

$$CD = q_{\alpha} \sqrt{\frac{k(k+1)}{6N}}$$

If the CD is greater than the values presented in the following Table, we can conclude that both algorithms have differences in performance:

#classifiers	2	3	4	5	6	7	8	9	10
$q_{0.05}$	1.960	2.241	2.394	2.498	2.576	2.638	2.690	2.724	2.773
$q_{0.10}$	1.645	1.960	2.128	2.241	2.326	2.394	2.450	2.498	2.539

⁽b) Critical values for the two-tailed Bonferroni-Dunn test; the number of classifiers include the control classifier.

Considering the example of the four versions of C4.5, we have (C4.5cf,m as control):

	C4.5	C4.5m	C4.5cf	C4.5cf,m
Average Rank	3.143	2.000	2.893	1.964

$$CD_{\alpha=0.05} = 2.394 \sqrt{\frac{4.5}{6.14}} = 1.16$$

$$CD_{\alpha=0.1}=2.128\sqrt{\frac{4.5}{6.14}}=1.038$$

With α =0.05, C4.5cf,m performs better than C4.5.

With α =0.1, C4.5cf,m also performs better than C4.5.

However, a more general way to obtain the differences among algorithms is to obtain a statistic that follow a normal distribution. The test statistics for comparing the i-th algorithm with the j-th algorithm is computed by:

$$z = (R_i - R_j) / \sqrt{\frac{k(k+1)}{6N}}$$

The z value is used to find the corresponding probability from the table of normal distribution, which is then compared with an appropriate α .

In Bonferroni-Dunn, α is always divided by (k-1) independently of the comparison, following a very conservative behavior. For this reason other procedures such as Holm's or Hochberg's are preferred.

Holm's method: We dispose of a test that sequentially checks the hypothesis ordered according to their significance. We will denote the p values ordered: $p_1 \le p_2 \le ... \le p_{k-1}$.

Holm's method compares each p_i with $\alpha/(k-i)$ starting from the most significant p value. If p_1 Is below than $\alpha/(k-1)$, the corresponding hypothesis is rejected and it leaves us to compare p_2 with $\alpha/(k-2)$. If the second hypothesis is rejected, we continue with the process. As soon as a certain hypothesis can not be rejected, all the remaining hypothesis are maintained as accepted.

The value of z is used for finding the corresponding probability from the table of the nomal distribution, which is compared with the corresponding value of α . (Table B.2. Zar – Proportions of the Normal Curve (One-tailed), App 17)

For Multiple Comparisons involving a Control Method

Holm's method: SE =
$$\sqrt{(4.5/6.14)} = 0.488$$
. $z = (R_i - R_j) / \sqrt{\frac{k(k+1)}{6N}}$.

p-values are

$$0.019 (C4.5+m)$$

0.016 (C4.5+m+cf).

i	classifier	$z = (R_0 - R_i)/SE$	p	α/i
1	C4.5+m+cf	(3.143 - 1.964)/0.488 = 2.416	0.016	0.017
2	C4.5+m	(3.143 - 2.000)/0.488 = 2.342	0.019	0.025
3	C4.5+cf	(3.143 - 2.893)/0.488 = 0.512	0.607	0.050

The first one is rejected (0.016 < 0.017)

The second one is rejected (0.019 < 0.025),

The third one can not be rejected (0.607 > 0.05)

Hochberg's method: It is a step-up procedure that works in the opposite direction to Holm's method, comparing the largest p value with α , the next largest with $\alpha/2$ and so forth until it encounters a hypothesis it can reject. All hypotheses with smaller p values are then rejected as well.

Hochberg's method is more powerful than Holm's although it may under some circumstances exceed the family-wise error.

Statistical Analysis of Experiments in Data Mining and Computational Intelligence

Basic Non-Parametric Tests and Case Studies

- For Pairwise Comparisons
- For Multiple Comparisons involving a Control Method
- Data Mining: Neural Networks and Genetic Learning
- **Evolutionary Algorithms: CEC'05 Special Session of Parameter Optimization**

Data Mining: Neural Networks and Genetic Learning

Wilcoxon Signed-Ranks Test for Paired Samples

Wilcoxon's test applied over the all possible comparisons between the algorithms in accuracy

Table 11 Wilcoxon test applied over the all possible comparisons between the five algorithms in classification rate

Comparison	Classification rate				
	R^+	R^-	p value		
Pitts-GIRLA-XCS	0.5	104.5	0.001		
Pitts-GIRLA-GASSIST-ADI	0	105	0.001		
Pitts-GIRLA-HIDER	1	104	0.001		
Pitts-GIRLA-CN2	6	99	0.004		
XCS-GASSIST-ADI	89	16	0.022		
XCS-HIDER	53	52	0.975		
XCS-CN2	78	27	0.109		
GASSIST-ADI-HIDER	20	85	0.041		
GASSIST-ADI-CN2	52	53	0.975		
HIDER-CN2	100	5	0.003		

We stress in **bold** the winner algorithm in each row when

the *p-value* associated is below 0.05

Data Mining: Neural Networks and Genetic Learning

Wilcoxon Signed-Ranks Test for Paired Samples

Wilcoxon's test applied over the all possible comparisons between the algorithms in kappa rate

Table 12 Wilcoxon test applied over the all possible comparisons between the five algorithms in kappa

Comparison	Cohen's kappa				
	R^+	R^-	p value		
Pitts-GIRLA-XCS	0.5	104.5	0.001		
Pitts-GIRLA-GASSIST-ADI	0	105	0.001		
Pitts-GIRLA-HIDER	0	105	0.001		
Pitts-GIRLA-CN2	10	95	0.008		
XCS-GASSIST-ADI	74	31	0.177		
XCS-HIDER	51	54	0.925		
XCS-CN2	78	27	0.109		
GASSIST-ADI-HIDER	28	77	0.124		
GASSIST-ADI-CN2	60	45	0.638		
HIDER-CN2	96	9	0.006		

We stress in **bold** the winner algorithm in each row when

the *p-value* associated is below 0.05

Data Mining: Neural Networks and Genetic Learning

Results of applying Friedman's and Iman-Davenport's test with level of significance $\alpha \le 0.05$ to the GBMLs

Table 13 Results of the Friedman and Iman–Davenport tests ($\alpha = 0.05$)

	Friedman Value	Value in χ ²	p value	Iman-Davenport Value	Value in F_F	p value
Classification rate	28.957	9.487	< 0.0001	13.920	2.55	< 0.0001
Cohen's kappa	26.729	9.487	< 0.0001	11.871	2.55	< 0.0001

- The statistics of Friedman and Iman-Davenport are clearly greater than their associated critical values
 - There are significant differences among the observed results
- Next step: apply post-hoc test and find what algorithms partners' average results are dissimilar

Data Mining: Neural Networks and Genetic Learning

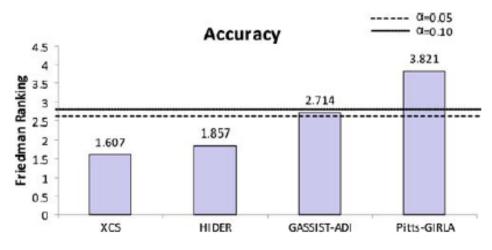


Fig. 3 Bonferroni-Dunn graphic for classification rate

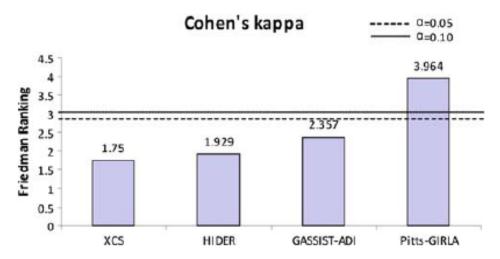


Fig. 4 Bonferroni-Dunn graphic for kappa

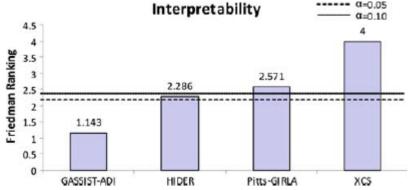


Fig. 5 Bonferroni-Dunn graphic measuring interpretability

Data Mining: Neural Networks and Genetic Learning

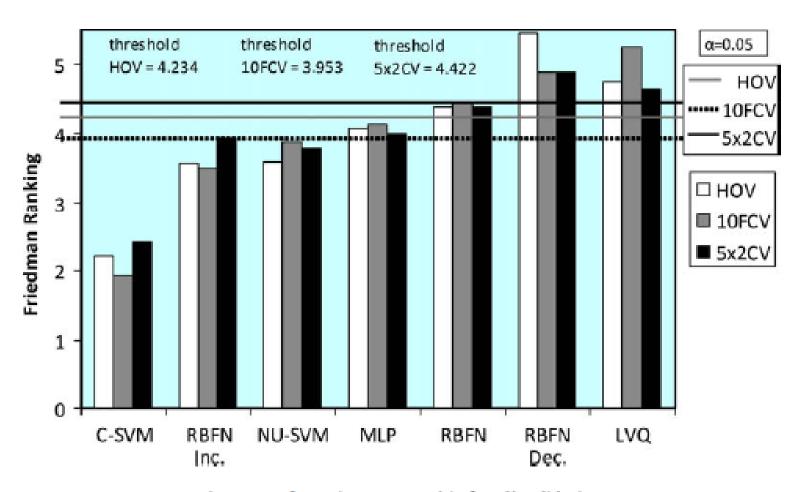


Fig. 5. Bonferroni-Dunn graphic for all validations.

Statistical Analysis of Experiments in Data Mining and Computational Intelligence

Basic Non-Parametric Tests and Case Studies

- **■** For Pairwise Comparisons
- For Multiple Comparisons involving a Control Method
- Data Mining: Neural Networks and Genetic Learning
- Evolutionary Algorithms: CEC'05 Special Session of Parameter Optimization

Evolutionary Algorithms: CEC'2005 Special Session of Parameter Optimization

TABLE XVI
WILCOXON TEST FOR ALL FUNCTIONS (F1-F25)

alg.	R^+	R^-	Hyp. α	Hyp. α	Hyp. α	Hyp. α
			0.01	0.02	0.05	0.1
BLX-GL50	289.5	35.5	R	R	R	R
BLX-MA	295.5	29.5	R	R	R	R
COEVO	301.0	24.0	R	R	R	R
DE	262.5	62.5	R	R	R	R
DMS-L-PSO	199.0	126.0	A	A	A	A
EDA	284.5	40.5	R	R	R	R
K-PCX	269.0	56.0	R	R	R	R
L-CMA-ES	273.0	52.0	R	R	R	R
L-SADE	209.0	116.0	A	A	A	A
SPC-PNX	305.5	19.5	R	R	R	R

G-CMAES versus the remaining algorithms.

The critical values are: 68, 76, 89 and 100 (0.01, 0.02, 0.05, 0.1)

Evolutionary Algorithms: CEC'2005 Special Session of Parameter Optimization

G-CMA-ES vs.	R^+	R ⁻	<i>p</i> -value
BLX-GL50	289.5	35.5	0.001
BLX-MA	295.5	29.5	0.001
CoEVO	301.0	24.0	0.000
DE	262.5	62.5	0.009
DMS-L-PSO	199.0	126.0	0.357
EDA	284.5	40.5	0.001
K-PCX	269.0	56.0	0.004
L-CMA-ES	273.0	52.0	0.003
L-SaDE	209.0	116.0	0.259
SPC-PNX	305.5	19.5	0.000

G-CMAES versus the remaining algorithms.
P-value obtained through normal approximation

Evolutionary Algorithms: CEC'2005 Special Session of Parameter Optimization

Example on the use of Wilcoxon's test combined for multiple comparisons

$$p = P(Reject \ H_0|H_0 \ true) =$$

$$= 1 - P(Accept \ H_0|H_0 \ true) =$$

$$= 1 - P(Accept \ A_k = A_i, i = 1, ..., k - 1|H_0 \ true) =$$

$$= 1 - \prod_{i=1}^{k-1} P(Accept A_k = A_i|H_0 \ true) =$$

$$= 1 - \prod_{i=1}^{k-1} [1 - P(Reject \ A_k = A_i|H_0 \ true)] =$$

$$= 1 - \prod_{i=1}^{k-1} (1 - p_{H_i})$$

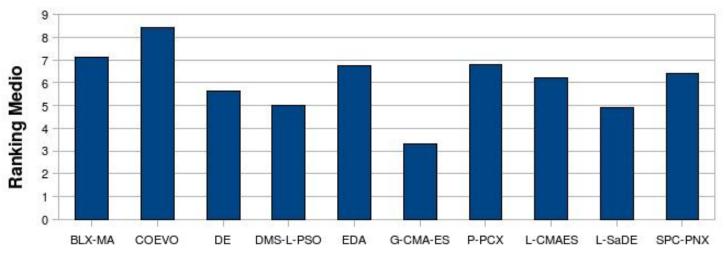
$$p = 1 - (1 - 0.001)(1 - 0.001)(1 - 0.000)(1 - 0.009)(1 - 0.357)$$
$$(1 - 0.001)(1 - 0.004)(1 - 0.003)(1 - 0.259)(1 - 0.000) = 0.467$$

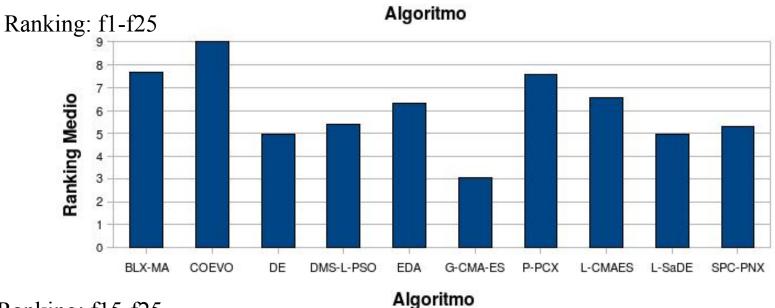
Table 7 Results of the Friedman and Iman-Davenport tests ($\alpha = 0.05$)

	Friedman value	Value in χ ²	<i>p</i> -value	Iman-Davenport value	Value in F_F	<i>p</i> -value
f15-f25	26.942	18.307	0.0027	3.244	1.930	0.0011
All	41.985	18.307	< 0.0001	4.844	1.875	< 0.0001

Ranking (f15–f25) Ranking (f1	
5.227	5.3
7.681	7.14
9.000	6.44
4.955	5.66
5.409	5.02
6.318	6.74
3.045	3.34
7.545	6.8
6.545	6.22
4.956	4.92
5.318	6.42
	5.227 7.681 9.000 4.955 5.409 6.318 3.045 7.545 6.545 4.956

Evolutionary Algorithms: CEC'2005 Special Session of Parameter Optimization





Ranking: f15-f25

97

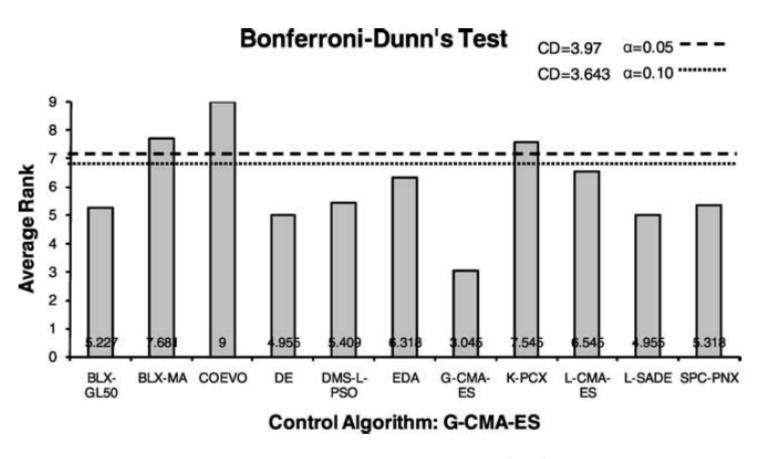


Fig. 6 Bonferroni-Dunn's graphic corresponding to the results for f15–f25

HOLM/HOCHBERG TABLE FOR FUNCTIONS F1-F25 (G-CMA-ES IS THE CONTROL ALGORITHM)

				4.	1.
ż	algorithm	z	p	α/i	α/i
				0.05	0.10
10	COEVO	5.43662	$5.43013 \cdot 10^{-8}$	0.00500	0.01000
9	BLX-MA	4.05081	$5.10399 \cdot 10^{-5}$	0.00556	0.01111
8	K-PCX	3.68837	$2.25693 \cdot 10^{-4}$	0.00625	0.01250
7	EDA	3.62441	$2.89619 \cdot 10^{-4}$	0.00714	0.01429
6	SPC-PNX	3.28329	0.00103	0.00833	0.01667
5	L-CMA-ES	3.07009	0.00214	0.01000	0.02000
4	DE	2.47313	0.01339	0.01250	0.02500
3	BLX-GL50	2.08947	0.03667	0.01667	0.03333
2	DMS-L-PSO	1.79089	0.07331	0.02500	0.05000
1	L-SADE	1.68429	0.09213	0.05000	0.10000

HOLM/HOCHBERG TABLE FOR FUNCTIONS F1-F25 (G-CMA-ES IS THE CONTROL ALGORITHM)

i	algorithm	z	p	α/i	α/i
			(0.05	0.10
10	COEVO	5.43662	$5.43013 \cdot 10^{-8}$	0.00500	0.01000
9	BLX-MA	4.05081	$5.10399 \cdot 10^{-5}$	0.00556	0.01111
8	K-PCX	3.68837	$2.25693 \cdot 10^{-4}$	0.00625	0.01250
7	EDA	3.62441	$2.89619 \cdot 10^{-4}$	0.00714	0.01429
6	SPC-PNX	3.28329	0.00103	0.00833	0.01667
5	L-CMA-ES	3.07009	0.00214	0.01000	0.02000
4	DE	2.47313	0.01339	0.01250	0.02500
3	BLX-GL50	2.08947	0.03667	0.01667	0.03333
2	DMS-L-PSO	1.79089	0.07331	0.02500	0.05000
1	L-SADE	1.68429	0.09213	0.05000	0.10000

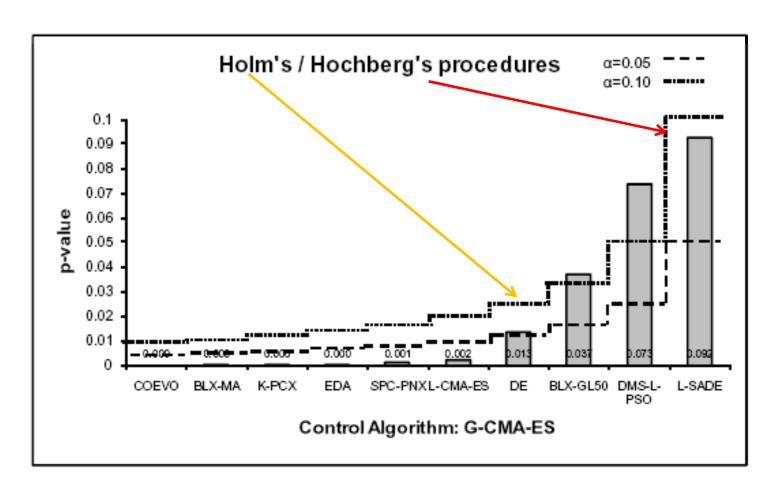


Fig. 11. Holm's/Hochberg's procedure for all functions (f1–f25).

Statistical Analysis of Experiments in Data Mining and Computational Intelligence

OUTLINE (II)

- Advanced non-parametric tests and case studies:
 - For Multiple Comparisons involving control method
 - Post-hoc Procedures
 - Adjusted p-values
 - Detecting all pairwise differences in a multiple comparison
- Lessons Learned
 - Considerations on the use of nonparametric tests
 - **Recommendations on the use of nonparametric tests**
 - **■** Frequent Questions
- Books of Interest and References
- Software

Statistical Analysis of Experiments in Data Mining and Computational Intelligence

Advanced non-parametric tests and case studies

- For Multiple Comparisons involving a control method
- Post-hoc Procedures
- Adjusted p-values
- Detecting all pairwise differences in a multiple comparison

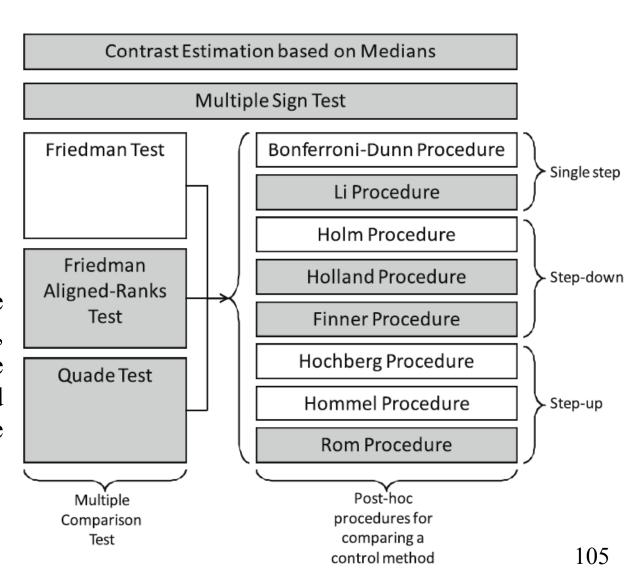
General Case Study used:

- □ 24 data sets from UCI and KEEL data-set
- Classifiers (from KEEL, standard parameters values):
 - PDFC
 - NNEP
 - □ IS-CHC + 1NN
 - FH-GBML
- 3 runs of 10fcv

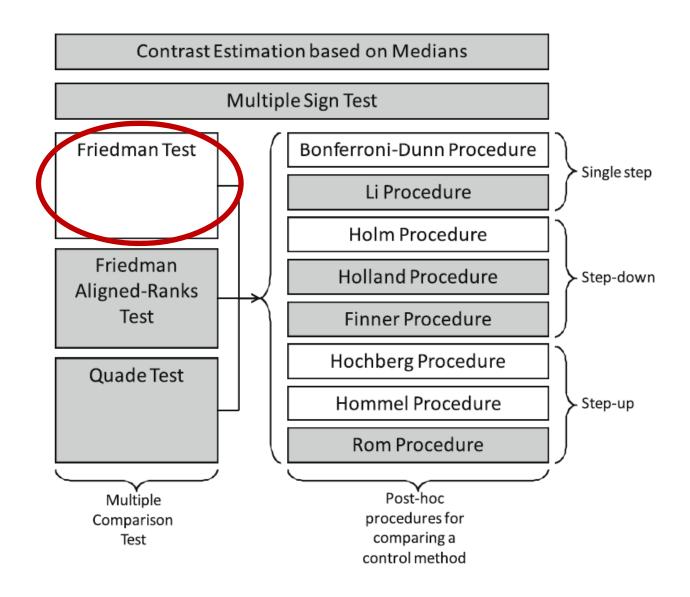
For Multiple Comparisons involving a Control Method

Multiple Comparison nonparametric procedures map.

In white are depicted the basic non-parametric test, whereas in grey are depicted more advanced tests which will be presented next.



For Multiple Comparisons involving a Control Method



For Multiple Comparisons involving a Control Method

•	1		
HT	ed	m	an

And

Iman-Davenport

(only showed for comparison purposes in this case study)

Dataset	PDFC	NNEP	IS-CHC+1NN	FH-GBML
adult	0,752 (4)	0,773 (3)	0,785 (2)	0,795(1)
breast	0,727(2)	0,748(1)	0,724(3)	0,713(4)
bupa	0,736(1)	0,716(2)	0,585 (4)	0,638 (3)
car	0,994(1)	0,861 (3)	0,880 (2)	0,791 (4)
cleveland	0,508 (4)	0,553(2)	0,575(1)	0,515(3)
contraceptive	0,535(2)	0,536(1)	0,513 (3)	0,471(4)
dermatology	0,967(1)	0.871(3)	0,954(2)	0,532(4)
ecoli	0,831(1)	0,807(3)	0,819(2)	0,768(4)
german	0,745 (1)	0,702(4)	0,719(2)	0,705(3)
glass	0,709(1)	0,572(4)	0,669 (2)	0,607 (3)
haberman	0,722(4)	0,728(2)	0,725(3)	0,732(1)
iris	0,967(1)	0,947(4)	0,953(3)	0,960(2)
lymphography	0,832(1)	0,752(3)	0,802(2)	0,691 (4)
mushrooms	0,998(1)	0,992(2)	0,482 (4)	0,910(3)
newthyroid	0,963 (1,5)	0,963 (1,5)	0,954(3)	0,926(4)
penbased	0,982(1)	0,953(2)	0,932(3)	0,630(4)
ring	0,978(1)	0,773(4)	0,834(3)	0,849(2)
satimage	0.854(1)	0,787(3)	0,841(2)	0,779(4)
shuttle	0,965 (3)	0,984(2)	0,995(1)	0,947(4)
spambase	0,924(1)	0,887(2)	0,861 (3)	0,804(4)
thyroid	0,929(3)	0,942(1)	0,931(2)	0,921(4)
vehicle	0,837(1)	0,643(2)	0,602 (3)	0,554 (4)
wine	0,972(1)	0,956(2)	0,944(3)	0,922(4)
wisconsin	0,958 (4)	0,959 (3)	0,964 (1,5)	0,964 (1,5)
Average ranking	1,771	2,479	2,479	3,271

For Multiple Comparisons involving a Control Method

Friedman's measure: 16.255

$$\chi_F^2 = \frac{12N}{k(k+1)} \left[\sum_j R_j^2 - \frac{k(k+1)^2}{4} \right]$$

Iman-Davenport's test:

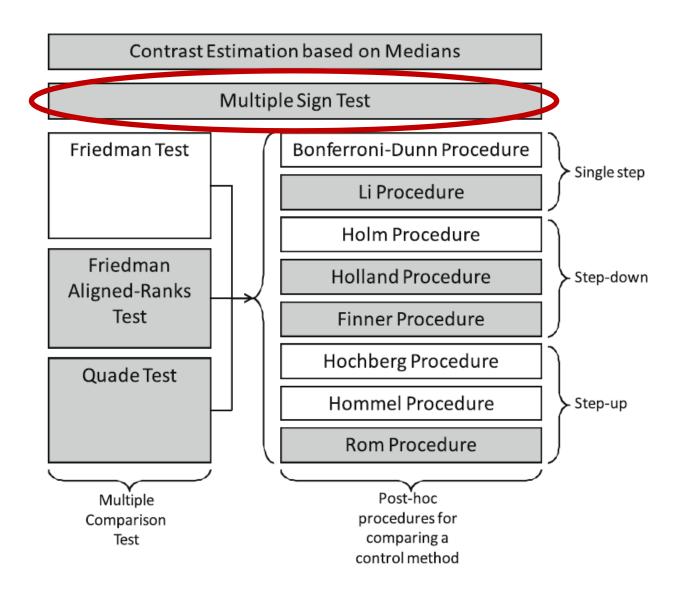
$$F_F = 6.691$$
, p-value for $F(3,3*23) = 0.000497$,

Therefore the null hypothesis is rejected.

$$F_F = \frac{(N-1)\chi_F^2}{N(k-1) - \chi_F^2}$$

(Friedman)
$$\chi_F^2 = \frac{12 \cdot 24}{4 \cdot 5} \left[(1.771^2 + 2.479^2 + 2.479^2 + 3.271^2) - \frac{4 \cdot 5^2}{4} \right] = 16.225$$
(Iman—Davenport)
$$F_F = \frac{23 \cdot 16.225}{24 \cdot 3 - 16.225} = 6.691$$

For Multiple Comparisons involving a Control Method



Advanced Non-Parametric Tests and Case Studies For Multiple Comparisons involving a Control Method

Multiple sign test: The following procedure, allows us to compare all of the other algorithms with a control labeled algorithm. The technique, an extension of the familiar sign test, carries out the following steps:

- 1. Represent by x_{i1} and x_{ij} the performances of the control and the jth classifier in the ith data set.
- 2. Compute the signed differences $d_{ij} = x_{ij} x_{i1}$. In other words, pair each performance with the control and, in each data set, subtract the control performance from the jth classifier.
- 3. Let r_j equal the number of differences, d_{ij} , that have the less frequently occurring sign (either positive or negative) within a pairing of an algorithm with the control.

Advanced Non-Parametric Tests and Case Studies For Multiple Comparisons involving a Control Method

- 4. Let M_1 be the median response of a sample of results of the control method and M_j be the median response of a sample of results of the jth algorithm. Apply one of the following decision rules:
 - For testing H_0 : $M_j \ge M_1$ against H_1 : $M_j < M_1$, reject H_0 if the number of plus signs is less than or equal to the critical value of R_j appearing in Table A.1 in Appendix A (Ref. below) for k 1 (number of algorithms excluding control), n and the chosen experimentwise error rate.
 - For testing H_0 : $M_j \le M_1$ against H_1 : $M_j Z M_1$, reject H_0 if the number of minus signs is less than or equal to the critical value of R_j appearing in Table A.1 in Appendix A for k 1, n and the chosen experimentwise error rate.

Source: S. García, A. Fernández, <u>J. Luengo</u>, <u>F. Herrera</u>, Advanced nonparametric tests for multiple comparisons in the design of experiments in computational intelligence and data mining: Experimental Analysis of Power. *Information Sciences 180 (2010) 2044–2064.* doi:10.1016/j.ins.2009.12.010.

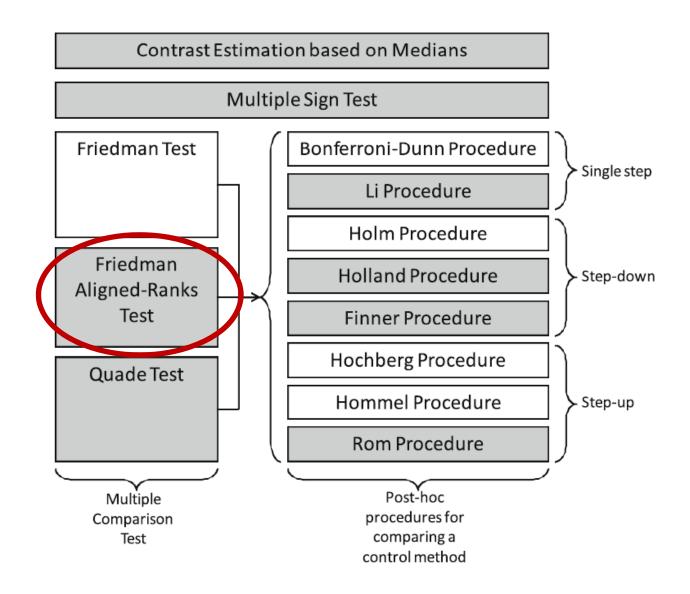
For Multiple Comparisons involving a Control Method

Multiple Sign

Test

Dataset	PDFC 1 (Control)	NNEP 2	IS-CHC+1NN 3	FH-GBML
adult	0,752	0,773 (+)	0,785 (+)	0.795 (+)
breast	0,727	0.748(+)	0,724 (-)	0,713 (-)
bupa	0,736	0,716 (-)	0,585 (-)	0,638 (-)
car	0,994	0,861 (-)	0,880 (-)	0,791 (-)
cleveland	0,508	0,553 (-)	0,575 (+)	0.515(+)
contraceptive	0,535	0.536(+)	0,513 (-)	0,471 (-)
dermatology	0,967	0,871 (-)	0,954 (-)	0,532 (-)
ecoli	0,831	0,807 (-)	0,819 (-)	0,768 (-)
german	0,745	0,702 (-)	0,719 (-)	0,705 (-)
glass	0,709	0,572 (-)	0,669 (-)	0,607 (-)
haberman	0,722	0,728 (+)	0,725 (+)	0,732 (+)
iris	0,967	0,947 (-)	0,953 (-)	0,960 (-)
lymphography	0,832	0,752 (-)	0,802 (-)	0,691 (-)
mushrooms	0,998	0,992 (-)	0,482 (-)	0,910 (-)
newthyroid	0,963	0,963 (=)	0,954 (-)	0,926 (-)
penbased	0,982	0,953 (-)	0,932 (-)	0,630 (-)
ring	0,978	0,773 (-)	0,834 (-)	0,849 (-)
satimage	0,854	0,787 (-)	0,841 (-)	0,779 (-)
shuttle	0,965	0,984 (+)	0,995 (+)	0,947 (-)
spambase	0,924	0,887 (-)	0,861 (-)	0,804 (-)
thyroid	0,929	0,942 (+)	0,931 (+)	0,921 (-)
vehicle	0,837	0,643 (-)	0,602 (-)	0,554 (-)
wine	0,972	0,956 (-)	0,944 (-)	0,922 (-)
wisconsin	0,958	0,959 (+)	0,964 (+)	0,964 (+)
Number of minus		16	18	20
Number of plus .		7	6	4
r _i		7	6	4
-1	•			

For Multiple Comparisons involving a Control Method



Advanced Non-Parametric Tests and Case Studies For Multiple Comparisons involving a Control Method

Aligned Ranks Friedman's test: a value of location is computed as the average performance achieved by all algorithms in each data set. Then, it calculates the difference between the performance obtained by an algorithm and the value of location. This step is repeated for algorithms and data sets. The resulting differences, called aligned observations, which keep their identities with respect to the data set and the combination of algorithms to which they belong, are then ranked from 1 to kn relative to each other. Then, the ranking scheme is the same as that employed by a multiple comparison procedure which employs independent samples; such as the Kruskal–Wallis test. The ranks assigned to the aligned observations are called aligned ranks.

$$T = \frac{(k-1)\left[\sum_{j=1}^{k} \widehat{R}_{,j}^{2} - (kn^{2}/4)(kn+1)^{2}\right]}{\{[kn(kn+1)(2kn+1)]/6\} - (1/k)\sum_{i=1}^{n} \widehat{R}_{i.}^{2}}$$

For Multiple Comparisons involving a Control Method

Friedman Aligned Ranks

Dataset	PDFC	NNEP	IS-CHC-1NN	FH-GBML	
adult	-0,024 (74)	-0,003 (56)	0,009 (39)	0,019 (30)	=
breast	-0,001 (51)	0,020 (29)	-0,004 (59)	-0,015 (68)	
bupa	0,068 (11)	0,047 (16)	-0,084 (90)	-0,031 (81)	
car	0,112(7)	-0,020 (72)	-0,002 (53)	-0,091 (92)	
cleveland	-0,030 (80)	0,016 (32)	0,037 (19)	-0,023 (73)	
contraceptive	0,022 (28)	0,022 (26)	-0,001 (50)	-0,043 (85)	
dermatology	0,136 (4)	0,040 (17)	0,123 (5)	-0,299 (95)	
ecoli	0,025 (24)	0,001 (48)	0,013 (33)	-0,038 (84)	
german	0,027 (22)	-0,016 (69)	0,001 (47)	-0,013 (67)	
glass	0,069 (10)	-0,068 (88)	0,030 (21)	-0,032 (82)	
haberman	-0,005 (61)	0,002 (46)	-0,002 (54)	0,005 (41)	
iris	0,010 (38)	-0,010 (66)	-0,003 (58)	0,003 (42)	
lymphography	0,063 (13)	-0,017 (71)	0,032 (20)	-0,078 (89)	
mushrooms	0,152(2)	0,146(3)	-0,363 (96)	0,065 (12)	
newthyroid	0,012 (34,5)	0,012(34,5)	0,002(45)	-0,026 (76)	
penbased	0,108 (8)	0,078 (9)	0,058 (14)	-0,244 (94)	
ring	0,120 (6)	-0,085 (91)	-0,025 (75)	-0,010 (65)	
satimage	0,038 (18)	-0,028 (79)	0,026 (23)	-0,036 (83)	
shuttle	-0,008 (62)	0,012 (36)	0,022 (27)	-0,026 (77)	
spambase	0,055 (15)	0,018 (31)	-0,008 (63)	-0,065 (87)	
thyroid	-0,001 (52)	0,011(37)	0,000 (49)	-0,010 (64)	
vehicle	0,178(1)	-0,016 (70)	-0,057 (86)	-0,105 (93)	
wine	0,024 (25)	0,007(40)	-0,004 (60)	-0,027 (78)	
wisconsin	-0,003 (57)	-0,002 (55)	0,003 (43,5)	0,003 (43,5)	
total	703,5	1121,5	1129,5	1701,5	115
average ranking	29,313	46,729	47,063	70,896	

For Multiple Comparisons involving a Control Method

Aligned Ranks Friedman's measure: 18.837

$$T = \frac{(k-1)\left[\sum_{j=1}^{k} \widehat{R}_{,j}^{2} - (kn^{2}/4)(kn+1)^{2}\right]}{\{[kn(kn+1)(2kn+1)]/6\} - (1/k)\sum_{i=1}^{n} \widehat{R}_{i.}^{2}}$$

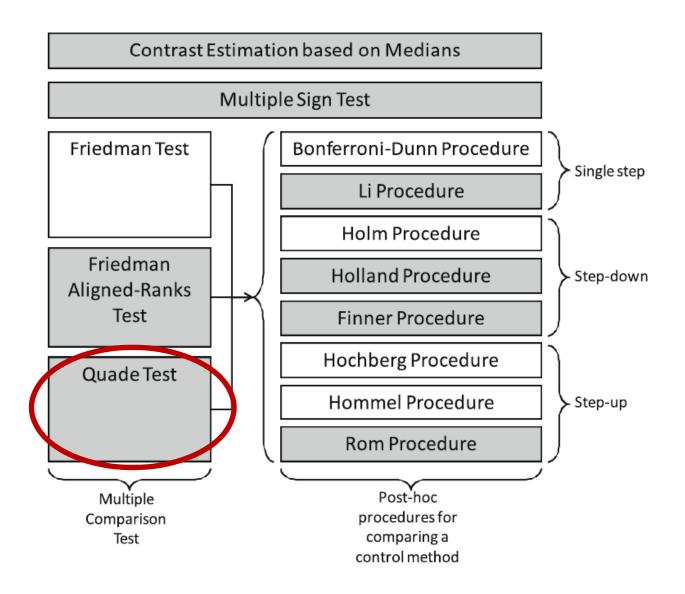
The p-value of Chi² with 3 degrees of freedom is 0.000296. Hypothesis rejected.

$$\sum_{j=1}^{k} \widehat{R}_{,j}^{2} = 703.5^{2} + 1121^{2} + 1129.5^{2} + 1701.5^{2} = 5,923,547$$

$$\sum_{i=1}^{n} \widehat{R}_{i.}^{2} = 199^{2} + 207^{2} + 198^{2} + \dots + 199^{2} = 926,830$$

$$T = \frac{(4-1)[5,923,547 - (4 \cdot 24^{2}/4)(4 \cdot 24 + 1)^{2}]}{\{[4 \cdot 24(4 \cdot 24 + 1)(2 \cdot 4 \cdot 24 + 1)]/6\} - (1/4) \cdot 926,830} = 18.837$$

For Multiple Comparisons involving a Control Method



Advanced Non-Parametric Tests and Case Studies For Multiple Comparisons involving a Control Method

Quade test: The Friedman test considers all data sets to be equal in terms of importance. An alternative to this could take into account the fact that some data sets are more difficult or the differences registered on the run of various algorithms over them are larger. The rankings computed on each data set could be scaled depending on the differences observed in the algorithms' performances.

The procedure starts finding the ranks in the same way as the Friedman test does. The next step requires the original values of performance of the classifiers. Ranks are assigned to the data sets themselves according to the size of the sample

range in each data set. The sample range within data set i is the difference between the largest and the smallest observations within that data set:

Range in data set :
$$i = \max_{j} \{x_{ij}\} - \min_{j} \{x_{ij}\}$$

$$S_{ij} = Q_i \left[r_i^j - \frac{k+1}{2} \right]$$

$$A_2 = n(n+1)(2n+1)(k)(k+1)(k-1)/72$$

$$B = \frac{1}{n} \sum_{j=1}^{k} S_j^2$$

The test statistic is

$$T_3 = \frac{(n-1)B}{A_2 - B}$$
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For Multiple Comparisons involving a Control Method

Quade

D 4 4	Sample	Ranking					
Dataset	Ranking	Q_i	PDFC	NNEP	IS-CHC-1NN	FH-GBML	
adult	0,043	8	0,752 (12)(32)	0,773 (4)(24)	0,785 (-4)(16)	0,795 (-12)(8)	_
breast	0,035	5	0,727 (-2,5)(10)	0,748 (-7,5)(5)	0,724 (2,5)(15)	0,713 (7,5)(20)	
bupa	0,151	18	0,736 (-27)(18)	0,716 (-9)(36)	0,585 (27)(72)	0,638 (9)(54)	
car	0,203	19	0,994 (-28,5)(19)	0,861 (9,5)(57)	0,880 (-9,5)(38)	0,791 (28,5)(76)	
cleveland	0,067	13	0,508 (19,5)(52)	0,553 (-6,5)(26)	0,575 (-19,5)(13)	0,515 (6,5)(39)	
contraceptive	0,065	12	0,535 (-6)(24)	0,536 (-18)(12)	0,513 (6)(36)	0,471 (18)(48)	
dermatology	0,436	23	0,967 (-34,5)(23)	0,871 (11,5)(69)	0,954 (-11,5)(46)	0,532 (34,5)(92)	
ecoli	0,063	11	0,831 (-16,5)(11)	0,807 (5,5)(33)	0,819 (-5,5)(22)	0,768 (16,5)(44)	
german	0,043	7	0,745 (-10,5)(7)	0,702 (10,5)(28)	0,719 (-3,5)(14)	0,705 (3,5)(21)	
glass	0,137	16	0,709 (-24)(16)	0,572 (24)(64)	0,669 (-8)(32)	0,607 (8)(48)	
haberman	0,010	2	0,722 (3)(8)	0,728 (-1)(4)	0,725 (1)(6)	0,732 (-3)(2)	
iris	0,020	3	0,967 (-4,5)(3)	0,947 (4,5)(12)	0,953 (1,5)(9)	0,960 (-1,5)(6)	
lymphography	0,141	17	0,832 (-25,5)(17)	0,752 (8,5)(51)	0,802 (-8,5)(34)	0,691 (25,5)(68)	
mushrooms	0,515	24	0,998 (-36)(24)	0,992 (-12)(48)	0,482 (36)(96)	0,910 (12)(72)	
newthyroid	0,038	6	0,963 (-6)(9)	0,963 (-6)(9)	0,954 (3)(18)	0,926 (9)(24)	
penbased	0,352	22	0,982 (-33)(22)	0,953 (-11)(44)	0,932 (11)(66)	0,630 (33)(88)	
ring	0,205	20	0,978 (-30)(20)	0,773 (30)(80)	0,834 (10)(60)	0,849 (-10)(40)	
satimage	0,075	14	0,854 (-21)(14)	0,787 (7)(42)	0,841 (-7)(28)	0,779 (21)(56)	
shuttle	0,048	9	0,965 (4,5)(27)	0,984 (-4,5)(18)	0,995 (-13,5)(9)	0,947 (13,5)(36)	
spambase	0,120	15	0,924 (-22,5)(15)	0,887 (-7,5)(30)	0,861 (7,5)(45)	0,804 (22,5)(60)	
thyroid	0,021	4	0,929 (2)(12)	0,942 (-6)(4)	0,931 (-2)(8)	0,921 (6)(16)	
vehicle	0,282	21	0,837 (-31,5)(21)	0,643 (-10,5)(42)	0,602 (10,5)(63)	0,554 (31,5)(84)	
wine	0,050	10	0,972 (-15)(10)	0,956 (-5)(20)	0,944 (5)(30)	0,922 (15)(40)	
wisconsin	0,006	1	0,958 (1,5)(4)	0,959 (0,5)(3)	0,964 (-1)(1,5)	0,964 (-1)(1,5)	
suma of rankings							_
S_{j}			-332	11	27,5	293,5	
rankings medios							119
$T_j = \frac{W_j}{n(n+1)/2}$			1,393	2,537	2,592	3,478	**/
2 11(11+1)/2	'		1				

For Multiple Comparisons involving a Control Method

Quade measure: 21.967

With four algorithms and 24 data sets, T₃ is distributed according to the F distribution with 4-1=3 and (4-1)*(24-1)=69 degrees of freedom. The p-value computed by using the F distribution is 0.000000000429, so the null hypothesis is rejected at a high level of significance.

$$A_2 = n(n+1)(2n+1)(k)(k+1)(k-1)/72$$

$$B = \frac{1}{n} \sum_{j=1}^{k} S_j^2$$

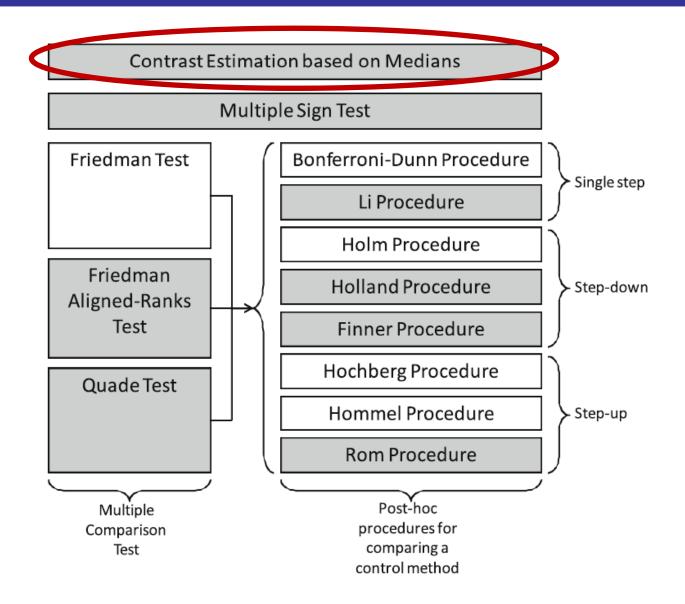
$$T_3 = \frac{(n-1)B}{A_2 - B}$$

$$A_2 = 24(24+1)(2 \cdot 24+1)4(4+1)(4-1)/72 = 24,500$$

$$B = \frac{1}{24}[(-332)^2 + 11^2 + 27.5^2 + 293.5^2] = 4068.479$$

$$T_3 = \frac{23 \cdot 4068.479}{24,500 - 4068.479} = 21.967$$

For Multiple Comparisons involving a Control Method



Advanced Non-Parametric Tests and Case Studies For Multiple Comparisons involving a Control Method

Contrast Estimation based on medians: Using the data resulting from the run of various classifiers over multiple data sets in an experiment, the researcher could be interested in the estimation of the difference between two classifiers' performance.

A procedure for this purpose assumes that the expected differences between performances of algorithms are the same across data sets. We assume that the performance is reflected by the magnitudes of the differences between the performances of the algorithms.

Consequently, we are interested in estimating the contrast between medians of samples of results considering all pairwise comparisons. It obtains a quantitative difference computed through medians between two algorithms over multiple data sets, but the value obtained will change when using other data sets in the experiment.

Advanced Non-Parametric Tests and Case Studies For Multiple Comparisons involving a Control Method

Contrast Estimation Based on Means procedure:

- 1. Compute the difference between every pair of k algorithms in each of the n data set: $D_{i(uv)} = X_{iu} X_{iv}$, only when u < v.
- 2. Compute the median of each set of differences Z_{uv} . It is the unadjusted estimator of $M_u M_v$. Sice $Z_{vu} = Z_{uv}$, we have only to calculate the cases u < v. $Z_{uu} = 0$.
- 3. Compute the mean of each set of unadjusted medians having the same first subscript m_u :

$$m_u = \frac{\sum_{j=1}^k Z_{uj}}{k}, u = 1, ..., k$$

4. The estimator of $M_u - M_v$ is $m_u - m_v$

For Multiple Comparisons involving a Control Method

Contrast Estimation based

on Medians

Dataset	$D_{i(12)}$	$D_{i(13)}$	$D_{i(14)}$	$D_{i(23)}$	$D_{i(24)}$	$D_{i(34)}$
adult*	-0,021	-0,033	-0,043	-0,012	-0,022	-0,010
breast	-0,021	0,003	0,014	0,024	0,035	0,011
bupa	0,020	0,151	0,099	0,131	0,078	-0,053
car	0,133	0,114	0,203	-0,019	0,071	0,089
cleveland	-0.045	-0.067	-0,007	-0.021	0.039	0.060
contraceptive	-0,001	0,022	0,064	0,023	0,065	0,042
dermatology	0,096	0,014	0,436	-0,083	0,339	0,422
ecoli	0,024	0,012	0,063	-0,012	0,039	0,051
german	0,043	0,026	0,040	-0.017	-0,003	0,014
glass	0,137	0,040	0,101	-0,097	-0,036	0,062
haberman	-0,006	-0,003	-0,010	0,004	-0,003	-0,007
iris	0,020	0,013	0,007	-0,007	-0,013	-0,007
lymphography	0,080	0,031	0,141	-0,049	0,061	0,110
mushrooms*	0,006	0,515	0,087	0,509	0,081	-0,428
newthyroid	0,000	0,010	0,038	0,010	0,038	0,028
penbased*	0,029	0,049	0,352	0,020	0,323	0,302
ring*	0,205	0,145	0,130	-0,061	-0,076	-0,015
satimage*	0,067	0,012	0,075	-0,054	0,008	0,062
shuttle*	-0,019	-0,030	0,018	-0,011	0,038	0,048
spambase*	0,037	0,063	0,120	0,026	0,083	0,057
thyroid*	-0,013	-0,001	0,008	0,011	0,021	0,010
vehicle	0,194	0,235	0,282	0,041	0,089	0,047
wine	0,016	0,028	0,050	0,011	0,034	0,023
wisconsin	-0,001	-0,006	-0,006	-0,005	-0,005	0,000

For Multiple Comparisons involving a Control Method

$$m_1 = \frac{0 + 0.02 + 0.018 + 0.064}{4} = 0.026$$

$$m_2 = \frac{-0.02 + 0 + (-0.006) + 0.038}{4} = 0.003$$

Our estimate is $m_1 - m_2$:

0.023

Contrast Estimation based on medians among all the algorithms of the case study presented

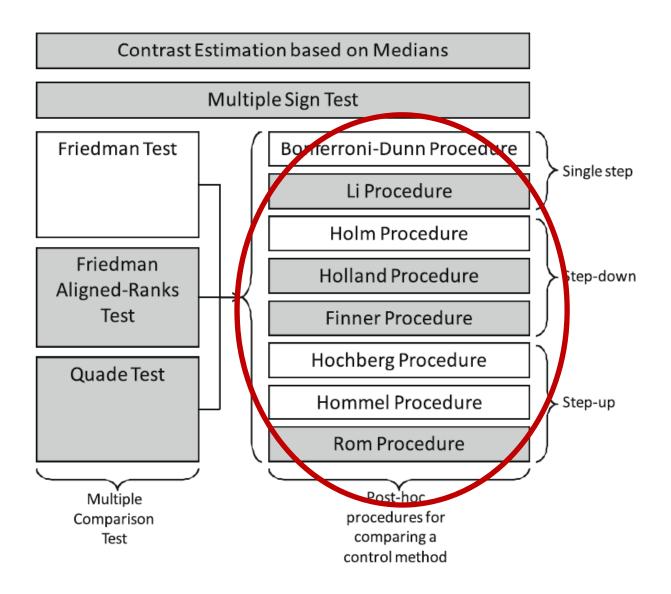
	PDFC	NNEP	IS-CHC+1NN	FH-GBML
PDFC	0.000	0.023	0.020	0.060
NNEP	-0.023	0.000	-0.003	0.037
IS-CHC+1NN	-0.020	0.003	0.000	0.040
FH-GBML	-0.060	-0.037	-0.040	0.000

Statistical Analysis of Experiments in Data Mining and Computational Intelligence

Advanced non-parametric tests and case studies

- For Multiple Comparisons involving a control method
- Post-hoc Procedures
- Adjusted p-values
- Detecting all pairwise differences in a multiple comparison

Post-hoc Procedures



Multiple Comparison tests are focused on the comparison between a control method, which is usually the proposed method, and a set of algorithms used in the empirical study. This set of comparisons is associated with a set or family of hypotheses, all of which are related to the control method. Any of the post hoc tests is suitable for application to nonparametric tests working over a family of hypotheses. The test statistic for comparing the ith algorithm and jth algorithm depends on the main nonparametric procedure used:

• Friedman
$$z = (R_i - R_j) / \sqrt{\frac{k(k+1)}{6n}}$$

• Friedman Aligned Ranks
$$z = (\widehat{R}_i - \widehat{R}_j) / \sqrt{\frac{k(n+1)}{6}}$$

•Quade
$$z = (T_i - T_j) / \sqrt{\frac{k(k+1)(2n+1)(k-1)}{18n(n+1)}}$$
 where $T_i = \frac{W_i}{n(n+1)/2}$, $T_j = \frac{W_j}{n(n+1)/2}$

REMEMBER: Three classical post-hoc procedures have been used in mutiple comparisons tests:

- **Bonferroni-Dunn:** controls the family-wise error rate by dividing a by the number of comparisons made (k-1).
- **Holm:** Step-down procedure that sequentially test the hypotheses ordered by their significance. We will denote the ordered p values by p_1 , p_2 , ..., so that $p_1 \le p_2 \le \ldots \le p_{k-1}$. It starts with the most significant p value. If p_1 is below $\alpha/(k-1)$, the corresponding hypothesis is rejected and we are allowed to compare p_2 with α /(k-2). If the second hypothesis is rejected, the test proceeds with the third, and so on.
- **Hochberg:** step-up procedure that works in the opposite direction, comparing the largest p value with α , the next largest with $\alpha/2$ and so forth until it encounters a hypothesis it can reject

Hommel: is more complicated to compute and understand. First, we need to find the largest j for which $p_{n-j+k} > k\alpha/j$ for all k = 1, ..., j. If no such j exists, we can reject all hypotheses, otherwise we reject all for which $p_i \le \alpha/j$.

Holland: it also adjusts the value of a in a step-down manner, as Holm's method does. It rejects H_1 to H_{i-1} if i is the smallest integer so that $p_i > 1 - (1 - \alpha)^{k-i}$.

Finner: it also adjusts the value of a in a step-down manner, as Holm's or Holland's method do. It rejects H_1 to H_{i-1} if i is the smallest integer so that $p_i > 1 - (1 - \alpha)^{(k-1)/i}$.

Rom: Rom developed a modification to Hochberg's procedure to increase its power. It works in exactly the same way as the Hochberg procedure, except that the a values are computed through the expression

$$\alpha_{k-i} = \left[\sum_{j=1}^{i-1} \alpha^j - \sum_{j=1}^{i-2} \binom{i}{k} \alpha_{k-1-j}^{i-j} \right] / i$$
where $\alpha_{k-1} = \alpha$ and $\alpha_{k-2} = \alpha/2$.

• Li - 2 steps rejection procedure:

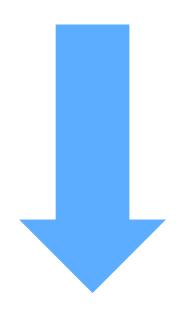
• Step 1: Reject all H_i if $p_{k-1} \le \alpha$. Otherwise, accept the hypothesis associated to p_{k-1} and got to step 2.

• Step 2: Reject any remaining H_i with $p_i \le (1-p_{k-1})/(1-\alpha)\alpha$

Post-hoc Procedures

A set of post-hoc procedures:

- one-step:
 - Bonferroni-Dunn
- step-down:
 - Holm
 - Holland
 - Finner
- step-up:
 - Hochberg
 - Hommel
 - Rom
- two-step:
 - Li



They are more powerful according this direction

Source: S. García, A. Fernández, <u>J. Luengo</u>, <u>F. Herrera</u>, Advanced nonparametric tests for multiple comparisons in the design of experiments in computational intelligence and data mining:

Experimental Analysis of Power. *Information Sciences 180 (2010) 2044–2064.*doi:10.1016/j.ins.2009.12.010.

Statistical Analysis of Experiments in Data Mining and Computational Intelligence

Advanced non-parametric tests and case studies

- For Multiple Comparisons involving a control method
- Post-hoc Procedures
- Adjusted p-values
- Detecting all pairwise differences in a multiple comparison

In statistical hypothesis testing, the p-value is the probability of obtaining a result at least as extreme as the one that was actually observed, assuming that the null hypothesis is true.

The smallest level of significance that results in the rejection of the null hypothesis, the *p-value*, is a useful and interesting datum for many consumers of statistical analysis.

A *p-value* provides information about whether a statistical hypothesis test is significant or not, and it also indicates something about "how significant" the result is: The smaller the *p-value*, the stronger the evidence against the null hypothesis. Most important, it does this without committing to a particular level of significance.

One way to solve this problem is to report adjusted p-values (APVs) which take into account that multiple tests are conducted.

An APV can be compared directly with any chosen significance level α .

We recommend the use of APVs due to the fact that they provide more information in a statistical analysis.

- Indexes *i* and *j* each correspond to a concrete comparison or hypothesis in the family of hypotheses, according to an incremental order of their *p*-values. Index *i* always refers to the hypothesis in question whose APV is being computed and index *j* refers to another hypothesis in the family.
- p_i is the p-value obtained for the jth hypothesis.
- *k* is the number of classifiers being compared.

APVs for each post-hoc procedure:

• one-step:

• Bonferroni-Dunn

APV_i: min{v; 1}, where $v = (k-1)p_i$.

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• step-down:

• Holm APV_i : min $\{v; 1\}$, where $v = \max\{(k-j)p_j : 1 \le j \le i\}$.

• Holland APV_i : $\min\{v; 1\}$, where $v = \max\{1 - (1 - p_j)^{k-j} : 1 \le j \le i\}$

• Finner APV_i : $\min\{v; 1\}$, where $v = \max\{1 - (1 - p_i)^{(k-1)/j} : 1 \le j \le i\}$

• step-up:

• Hochberg APV_i : $\max\{(k-j)p_j:(k-1)\geqslant j\geqslant i\}$

• Hommel (very difficult to compute, next slide)

• Rom $APV_i : \max\{(r_{k-j})p_j : (k-1) \ge j \ge i\}$

• two-step:

• Li $APV_i : p_i/(p_i + 1 - p_{k-1})$

```
1. Set APV_i = p_i for all i.
2. For each j = k - 1, k - 2, ..., 2 (in that order)
   3. Let B = \emptyset.
   4. For each i, i > (k - 1 - j)
      5. Compute value c_i = (j \cdot p_i)/(j+i-k+1).
      6. B = B \cup c_i.
   7. End for
   8. Find the smallest c_i value in B; call it c_{min}.
   9. If APV_i < c_{min}, then APV_i = c_{min}.
    10. For each i, i \le (k - 1 - j)
      11. Let c_i = min(c_{min}, j \cdot p_i).
      12. If APV_i < c_i, then APV_i = c_i.
    13. End for
```

Fig. 2. Algorithm for calculating APVs based on Hommel's procedure.

APVs in CEC'2005 Case Study

Table 10: p-values on functions f1-f25 (G-CMA-ES is the control algorithm)

			\	_	,
G-CMA-ES vs.	z	unadjusted p	Bonferroni-Dunn p	Holm p	Hochberg p
CoEVO	5.43662	$5.43013 \cdot 10^{-8}$	$5.43013 \cdot 10^{-7}$	$5.43013 \cdot 10^{-7}$	$5.43013 \cdot 10^{-7}$
BLX-MA	4.05081	$5.10399 \cdot 10^{-5}$	$5.10399 \cdot 10^{-4}$	$4.59359 \cdot 10^{-4}$	$4.59359 \cdot 10^{-4}$
K-PCX	3.68837	$2.25693 \cdot 10^{-4}$	0.002257	0.001806	0.001806
EDA	3.62441	$2.89619 \cdot 10^{-4}$	0.0028961	0.002027	0.002027
SPC-PNX	3.28329	0.00103	0.0103	0.00618	0.00618
L-CMA-ES	3.07009	0.00214	0.0214	0.0107	0.0107
$_{ m DE}$	2.47313	0.01339	0.1339	0.05356	0.05356
BLX-GL50	2.08947	0.03667	0.3667	0.11	0.09213
DMS-L-PSO	1.79089	0.07331	0.7331	0.14662	0.09213
L-SaDE	1.68429	0.09213	0.9213	0.14662	0.09213

In practice, Hochberg's method is more powerful than Holm's one (but this difference is rather small), in this the results are in favour of Hochberg's method.

APVs in GBMLs Case Study

Adjusted p -
values for the
comparison of
the control
algorithm in
each measure
with the
remaining
algorithms

i	Algorithm	Unadjusted p	$p_{ m Bonf}$	$p_{ m Holm}$	p_{Hoch}
Classifica	ation rate (XCS is the control))			-
1	Pitts-GIRLA	1.745×10^{-6}	6.980×10^{-6}	6.980×10^{-6}	6.980×10^{-6}
2	CN2	0.01428	0.05711	0.04283	0.04283
3	GASSIST-ADI	0.02702	0.10810	0.05405	0.05405
4	HIDER	0.67571	1.00000	0.67571	0.67571
Cohen's	kappa (XCS is the control)				
1	Pitts-GIRLA	5.576×10^{-6}	2.230×10^{-5}	2.230×10^{-5}	2.230×10^{-5}
2	CN2	0.01977	0.07908	0.05931	0.05931
3	GASSIST-ADI	0.13517	0.54067	0.27033	0.27033
4	HIDER	0.76509	1.00000	0.76509	0.76509

- If the adjusted p for each method is lower than the desired level of confidence α (0.05 in our case), the algorithms are worse from bottom to top (stress in bold for 0.05)
- In practice, Hochberg's method is more powerful than Holm's one (but this difference is rather small), in this our study the results are the same. 140

APVs in ANNs Case Study

Table 17 Adjusted *p*-values in 10FCV (*C*-SVM is the control).

i	Algorithm	Unadjusted p	p_{Bonf}	$p_{ m Holm}$	p_{Hoch}
1	LVQ	$1.443 \cdot 10^{-5}$	$8.663 \cdot 10^{-5}$	$8.663 \cdot 10^{-5}$	$8.663 \cdot 10^{-5}$
2	RBFN Decremental	$1.2 \cdot 10^{-4}$	$7.201 \cdot 10^{-4}$	$6.001 \cdot 10^{-4}$	$6.001 \cdot 10^{-4}$
3	RBFN	0.00106	0.00638	0.00425	0.00425
4	MLP	0.00418	0.02509	0.01255	0.01255
5	NU-SVM	0.01119	0.06713	0.02238	0.02238
6	RBFN Inc.	0.04078	0.24466	0.04078	0.04078

APVs for all post-hoc procedures in Friedman PDFC is the control

i Algorithm	1 FH-GBML	2 IS-CHC + 1NN	3 NNEP
Unadjusted p p _{Bonf} p _{Holm} p _{Hoch} p _{Homm} p _{Holl} p _{Rom} p _{Finn}	5.69941×10^{-5} 1.70982×10^{-4} 1.70982×10^{-4} 1.70982×10^{-4} 1.70982×10^{-4} 1.70973×10^{-4} 1.70982×10^{-4} 1.70982×10^{-4}	0.05735 0.17204 0.11469 0.05735 0.05735 0.11141 0.05735 0.08477	0.05735 0.17204 0.11469 0.05735 0.05735 0.11141 0.05735 0.08477
p_{Li}	6.04577×10^{-4}	0.05735	0.05735

APVs for all post-hoc procedures in Friedman Aligned Ranks PDFC is the control

i Algorithm	1 FH-GBML	2 IS-CHC + 1NN	3 NNEP
Unadjusted p p _{Bonf} p _{Holm} p _{Hoch} p _{Homm} p _{Holl} p _{Rom} p _{Finn}	2.32777×10^{-7} 6.98332×10^{-7}	0.02729 0.08188 0.05459 0.03032 0.03032 0.05384 0.03032 0.04066	0.03032 0.09097 0.05459 0.03032 0.03032 0.05384 0.03032 0.04066
p_{Li}	2.40057×10^{-7}	0.02738	0.03032

APVs for all post-hoc procedures in Quade

PDFC is the control

Algorithm FH-0	GBML	2 IS-CHC + 1NN	3 NNEP
p_{Holm} 1.93 p_{Hoch} 1.93 p_{Homm} 1.93 p_{Holm} 1.93 p_{Holl} 1.93 p_{Rom} 1.93 p_{Finn} 1.93	747×10^{-4} 124×10^{-4}	0.06490 0.04326 0.02843 0.02843 0.04280 0.02843 0.03227	0.02843 0.08528 0.04326 0.02843 0.02843 0.04280 0.02843 0.03227 0.02843

Statistical Analysis of Experiments in Data Mining and Computational Intelligence

Advanced non-parametric tests and case studies

- For Multiple Comparisons involving a control method
- Post-hoc Procedures
- Adjusted p-values
- Detecting all pairwise differences in a multiple comparison

Detecting all pairwise differences in a multiple comparison:

Until now, we have studied the techniques for multiple comparison using a control method. But, under some circumstances, it would be interesting to conduct a test over all possible comparisons involved in the experimental study

It is the usual case in review papers. In these cases, the repetition of comparisons choosing different control classifiers may lose the control of the family-wise error.

The post-hoc procedures need to control the FWER under more restrictive corrections because the family of hypotheses is formed now for k(k-1)/2 comparisons instead of (k-1).

Remember (Friedman):

A set of pairwise comparisons can be associated with a set or family of hypotheses. Any of the post-hoc tests which can be applied to non-parametric tests work over a family of hypotheses.

The test statistics for comparing the i-th and j-th classifier is

$$z = \frac{(R_i - R_j)}{\sqrt{\frac{k(k+1)}{6N}}}$$

The z value is used to find the corresponding probability (p-value) from the table of normal distribution, which is then compared with an appropriate level of significance a (Table A1 in Sheskin, 2003)

REMEMBER: Two classical post-hoc procedures have been used in mutiple comparisons tests and also valid in n x n comparisons:

- Bonferroni-Dunn (Nemenyi in n x n comparisons): controls the family-wise error rate by dividing a by the number of comparisons made m = k(k-1)/2.
- **Holm:** Step-down procedure that sequentially test the hypotheses ordered by their significance. We will denote the ordered p values by $p_1, p_2, ...,$ so that $p_1 \le p_2 \le ... \le p_{k-1}$. It starts with the most significant p value. If p_1 is below $\alpha/(m-1)$, the corresponding hypothesis is rejected and we are allowed to compare p_2 with $\alpha/(m-2)$. If the second hypothesis is rejected, the test proceeds with the third, and so on.

• Hochberg, Hommel, Rom, Finner are also valid....

Logically Related Hypotheses:

The hypotheses being tested belonging to a family of all pairwise comparisons are logically interrelated so that not all combinations of true and false hypotheses are possible.

As a simple example of such a situation suppose that we want to test the three hypotheses of pairwise equality associated with the pairwise comparisons of three classifiers C_i ; i = 1,2,3. It is easily seen from the relations among the hypotheses that if any one of them is false, at least one other must be false. For example, if C_1 is different than C_2 , then it is not possible that C_1 has the same performance than C_3 and C_2 has the same performance than C_3 . C_3 must be different than C_1 or C_2 or both.

Shaffer's procedure: following Holm's step down method, at stage j, instead of rejecting H_i if $p_i \le \alpha / (m-i+1)$, reject H_i if $p_i \le \alpha / t_i$, where t_i is the maximum number of hypotheses which can be true given that any (i - 1) hypotheses are false.

It is a static procedure, that is, t_1 , ..., t_m are fully determined for the given hypotheses H_1 , ..., H_m , independent of the observed p-values. The possible numbers of true hypotheses, and thus the values of t_i can be obtained from the recursive formula

$$S(k) = \bigcup_{j=1}^{k} \left\{ {j \choose 2} + x : x \in S(k-j) \right\},$$

where S(k) is the set of possible numbers of true hypotheses with k classifiers being compared, $k \ge 2$, and $S(0) = S(1) = \{0\}$.

Definition 1 *An index set of hypotheses* $I \subseteq \{1,...,m\}$ *is called* exhaustive *if exactly all* H_j , $j \in I$, *could be true.*

Bergmann-Hommel's procedure: Reject all Hj with j not in A, where the acceptance set

$$A = \bigcup \{I : I \text{ exhaustive}, \min\{P_i : i \in I\} > \alpha/|I|\}$$

is the index set of null hypotheses which are retained.

For this procedure, one has to check for each subset I of {1,...,m} if I is exhaustive, which leads to intensive computation. Due to this fact, we will obtain a set, named E, which will contain all the possible exhaustive sets of hypotheses for a certain comparison. Once the E set is obtained, the hypotheses that do not belong to the A set are rejected.

Detecting all Pairwise differences in a Multiple Comparison

```
Function obtainExhaustive(C = \{c_1, c_2, ..., c_k\}: list of classifiers)
1. Let E = \emptyset
2. E = E \cup \{\text{set of all possible and distinct pairwise comparisons using } C\}
3. If E == \emptyset
   4. Return E
5. End if
6. For all possible divisions of C into two subsets C_1 and C_2, c_k \in C_2 and C_1 \neq \emptyset
   7. E_1 = obtainExhaustive(C_1)
   8. E_2 = obtainExhaustive(C_2)
   9. E = E \cup E_1
   10. E = E \cup E_2
   11. For each family of hypotheses e_1 of E_1
      12. For each family of hypotheses e_2 of E_2
        13. E = E \cup (e_1 \cup e_2)
      14. End for
   15. End for
End for
17. Return E
```

Figure 1: Algorithm for obtaining all exhaustive sets

Detecting all Pairwise differences in a Multiple Comparison

Case Study used:

- □ 30 data sets from UCI and KEEL data-set
- Classifiers (from KEEL, standard parameters values):
 - □ C4.5
 - □ 1NN
 - Naïve Bayes
 - Kernel
 - CN2
- Rankings computed by Friedman test

Detecting all Pairwise differences in a Multiple Comparison

	C4.5	1NN	NaiveB	Kernel	CN2
Average Rank	2.100	3.250	2.200	4.333	3.117

i	hypothesis	$z = (R_0 - R_i)/SE$	p	α_{NM}	α_{HM}	α_{SH}
1	C4.5 vs. Kernel	5.471	$4.487 \cdot 10^{-8}$	0.005	0.005	0.005
2	NaiveBayes vs. Kernel	5.226	$1.736 \cdot 10^{-7}$	0.005	0.0055	0.0083
3	Kernel vs. CN2	2.98	0.0029	0.005	0.0063	0.0083
4	C4.5 vs. 1NN	2.817	0.0048	0.005	0.0071	0.0083
5	1NN vs. Kernel	2.654	0.008	0.005	0.0083	0.0083
6	1NN vs. NaiveBayes	2.572	0.0101	0.005	0.01	0.0125
7	C4.5 vs. CN2	2.49	0.0128	0.005	0.0125	0.0125
8	NaiveBayes vs. CN2	2.245	0.0247	0.005	0.0167	0.0167
9	1NN vs. CN2	0.327	0.744	0.005	0.025	0.025
10	C4.5 vs. NaiveBayes	0.245	0.8065	0.005	0.05	0.05

Table 3: Family of hypotheses ordered by p-value and adjusting of α by Nemenyi (NM), Holm (HM) and Shaffer (SH) procedures, considering an initial $\alpha=0.05$ 154

Detecting all Pairwise differences in a Multiple Comparison

Size 1	Size 2	Size 3	Size 4	$Size \ge 6$
(12)	(12,34)	(12,13,23)	(12,13,23,45)	(12,13,14,15,23,24,25,34,35,45)
(13)	(13,24)	(12,14,24)	(12,14,24,35)	(12,13,14,23,24,34)
(23)	(14,23)	(13,14,34)	(12, 34, 35, 45)	(12,13,15,23,25,35)
(14)	(12,35)	(23, 24, 34)	(13,14,25,34)	(12,14,15,24,25,45)
(24)	(13,25)	(12,15,25)	(13,15,24,35)	(13,14,15,34,35,45)
(34)	(15,23)	(13, 15, 35)	(13,24,25,45)	(23,24,25,34,35,45)
(15)	(12,45)	(23, 25, 35)	(14,15,23,45)	
(25)	(13,45)	(14, 15, 45)	(14,23,25,35)	
(35)	(23,45)	(24, 25, 45)	(15,23,24,34)	
(45)	(14,25)	(34, 35, 45)		
	(15,24)			
	(14,35)			
	(24,35)			

Exhaustive sets obtained for the case study. Those belonging to the *Acceptance* set (*A*) are typed in bold.

Case Study used:

- \square Nemenyi's test rejects the hypotheses [1–4] since the corresponding p-values are smaller than the adjusted α 's.
- ☐ Holm's procedure rejects the hypotheses [1–5].
- □ Shaffer's static procedure rejects the hypotheses [1–6].
- Bergmann-Hommel's dynamic procedure first obtains the exhaustive index set of hypotheses. It obtains 51 index sets. We can see them in the previous slide. From the index sets, it computes the A set. It rejects all hypotheses H_j with j not in A, so it rejects the hypotheses [1–8].

- m is the number of possible comparisons in an all pairwise comparisons design; that is, $m = \frac{k \cdot (k-1)}{2}$.
- t_j is the maximum number of hypotheses which can be true given that any (j-1) hypotheses are false

APVs for each post-hoc procedure:

• one-step:

• Nemenyi

 APV_i : $min\{v;1\}$, where $v=m \cdot p_i$.

• step-down:

• Holm

 APV_i : $min\{v, 1\}$, where $v = max\{(m-j+1)p_j : 1 \le j \le i\}$.

• Shaffer

 APV_i : $min\{v;1\}$, where $v = max\{t_j p_j : 1 \le j \le i\}$.

• Bergmann-Hommel

 APV_i : $min\{v;1\}$, where $v = max\{|I| \cdot min\{p_j, j \in I\} : I \text{ exhaustive}, i \in I\}$.

i	hypothesis	p_i	APV_{NM}	APV_{HM}	APV_{SH}	APV_{BH}
1	C4.5 vs .Kernel	$4.487 \cdot 10^{-8}$	$4.487 \cdot 10^{-7}$	$4.487 \cdot 10^{-7}$	$4.487 \cdot 10^{-7}$	$4.487 \cdot 10^{-7}$
2	NaiveBayes vs .Kernel	$1.736 \cdot 10^{-7}$	$1.736 \cdot 10^{-6}$	$1.563 \cdot 10^{-6}$	$1.042 \cdot 10^{-6}$	$1.042 \cdot 10^{-6}$
3	Kernel vs .CN2	0.0029	0.0288	0.023	0.0173	0.0115
4	C4.5 vs .1NN	0.0048	0.0485	0.0339	0.0291	0.0291
5	1NN vs .Kernel	0.008	0.0796	0.0478	0.0478	0.0319
6	1NN vs .NaiveBayes	0.0101	0.1011	0.0506	0.0478	0.0319
7	C4.5 vs .CN2	0.0128	0.1276	0.0511	0.0511	0.0383
8	NaiveBayes vs .CN2	0.0247	0.2474	0.0742	0.0742	0.0383
9	1NN vs .CN2	0.744	1.0	1.0	1.0	1.0
10	C4.5 vs .NaiveBayes	0.8065	1.0	1.0	1.0	1.0

APVs obtained in the example by Nemenyi (NM), Holm (HM), Shaffer's static (SH) Bergmann-Hommel's dynamic (BH)

Statistical Analysis of Experiments in Data Mining and Computational Intelligence

OUTLINE (II)

- Advanced non-parametric tests and case studies:
 - **■** For Multiple Comparisons involving control method
 - Post-hoc Procedures
 - Adjusted p-values
 - Detecting all pairwise differences in a multiple comparison
- Lessons Learned
 - Considerations on the use of nonparametric tests
 - **Recommendations on the use of nonparametric tests**
 - Frequent Questions
- Books of Interest and References
- Software

Statistical Analysis of Experiments in Data Mining and Computational Intelligence

Lessons Learned

- Considerations on the use of non-parametric tests
- Recommendations on the use of non-parametric tests
- Frequent questions

On the use of non-parametric tests:

The need of using non-parametric tests given that the necessary conditions for using parametric tests are not verified.

Considerations on the Use of Non-Parametric Tests

Wilcoxon's test

□ Wilcoxon's test computes a ranking based on differences between functions independently, whereas Friedman and derivative procedures compute the ranking between algorithms.

☐ Wilcoxon's test is highly influenced by the number of case of study (functions, data sets ...). The N value determines the critical values to search in the statistical table.

It is highly influenced by outliers when N is below or equal to 11.

Considerations on the Use of Non-Parametric Tests

Multiple comparison (1)

☐ A multiple comparison must be carried out first by using a statistical method for testing the differences among the related samples means. Then to use a post-hoc statistical procedures.

☐ Holm's procedure is a very good test.

Hochberg's method can rejects more hypothesis than Holm's one.

Considerations on the Use of Non-Parametric Tests

Multiple comparison (2)

- An appropriate number of algorithms in contrast with an appropriate number of case problems are needed to be used in order to employ each type of test. The number of algorithms used in multiple comparisons procedures must be lower than the number of case problems
- Both, the Friedman Aligned Rank test and the Quade test, can be used under the same circumstances as the Friedman test. The differences in power between them are unknown, but we encourage the use of these tests when the number of algorithms to be compared is low.

What happens if I use a nonparametric test when the data is normal?

- It will work, but a parametric test would be more powerful, i.e., give a lower p value.
- If the data is not normal, then the nonparametric test is usually more powerful
- Always look at the data first, then decide what test to use.

Advantages of Nonparametric Tests

- Can treat data which are inherently in ranks as well as data whose seemingly numerical scores have the strength in ranks
- Easier to learn and apply than parametric tests (only one run for all cases of test)

If sample sizes as small as N=6 are used, there is no alternative to using a nonparametric test

Advantages of Nonparametric Tests

If we have a set of data sets/benchmark functions, we must apply a parametric test for each data set/benchmark function.

We only need to use a non-parametric test for comparing the algorithms on the whole set of benchmarks.

Design of Experiments in Data Mining/Computational Intelligence

They are not the objective of our talk, but they are two additional important questions:

- ☐ Benchmark functions/data sets ... are very important.
- **□** To compare with the state of the art is a necessity.

Statistical Analysis of Experiments in Data Mining and Computational Intelligence

Lessons Learned

- **■** Considerations on the use of non-parametric tests
- Recommendations on the use of non-parametric tests
- Frequent questions

Recommendations on the Use of Non-Parametric Tests

Wilcoxon's test

An appropriate number of algorithms in contrast with an appropriate number of case problems are needed to be used in order to employ each type of test. The number of algorithms used in multiple comparisons procedures must be lower than the number of case problems. The previous statement may not be true in Wilcoxon's test. The influence of the number of case problems used is more noticeable in multiple comparisons procedures than in Wilcoxon's test.

Recommendations on the Use of Non-Parametric Tests

Multiple comparison with a control (1)

Holm's procedure can always be considered better than Bonferroni-Dunn's one, because it appropriately controls the FWER and it is more powerful than the Bonferroni-Dunn's. We strongly recommend the use of Holm's method in a rigorous comparison.

☐ Hochberg's procedure is more powerful than Holm's. The differences reported between it and Holm's procedure are in practice rather small. We recommend the use of this test together with Holm's method

Recommendations on the Use of Non-Parametric Tests

Multiple comparison with a control (2)

- □ Holm, Hochberg, Finner and Li are the more recommended post-hoc test to be used due to their trade-off between simplicity and power.
- ☐ The power of the Li test is highly influenced by the first p-value of the family and when it is lower than 0.5, the test will perform very well.
- ☐ The choice of any of the statistical procedures for conducting an experimental analysis should be justified by the researcher. The use of the most powerful procedures does not imply that the results obtained by his/her proposal will be better

Recommendations on the Use of Non-Parametric Tests

Multiple comparison with a control (3)

An alternative to directly performing a comparison between a control algorithm and a set of algorithms is the Multiple Sign-test. We recommend its use when the differences reported by the control algorithm with respect to the rest of methods are very clear.
☐ The Contrast Estimation in nonparametric statistics is used for computing the real differences between two algorithms, considering the median measure the most important.
☐ Finally, we want to remark that the choice of any of the statistical procedures presented in this talk for conducting an experimental analysis should be justified by the researcher. The use of the most powerful procedures does not imply that the results obtained by his/her proposal will be better.

Recommendations on the Use of Non-Parametric Tests

All pairwise comparisons in multiple comparison

☐ We do not recommend the use of Nemenyi's test, because it is a very conservative procedure and many of the obvious differences may not be detected. □ Conducting the Shaffer static procedure means a not very significant increase of the difficulty with respect to the Holm procedure. ☐ Bergmann-Hommel's procedure is the best performing one, but it is also the most difficult to understand and computationally expensive. We recommend its usage when the situation requires so.

Statistical Analysis of Experiments in Data Mining and Computational Intelligence

Lessons Learned

- **■** Considerations on the use of non-parametric tests
- Recommendations on the use of non-parametric tests
- Frequent questions

Lessons Learned Frequent Questions

☐ Can we analyze any performance measure?

☐ With non-parametric statistic, any unitary performance measure (associated to an only algorithm) with a pre-defined range of output can be analyzed. This range could be unlimited, allowing us to analyze time resources as example.

Lessons LearnedFrequent Questions

☐ Can we compare deterministic algorithms with stochastic ones?

☐ They allow us to compare both types of algorithms because they can be applied in multi-domain comparisons, where the sample of results is composed by a result that relates an algorithm and a domain of aplication (problem, function, data-set, ...)

Frequent Questions

☐ How the average results should be obtained from each algorithm?

This question does not concern to the use of non-parametric statistics, due to the fact that these tests require a result for each pair algorithm-domain. The obtaining of such result must be according to a standard procedure followed by all the algorithms in the comparison, such the case of validation techniques. Average results from various runs (at least 3) must be used for stochastic algorithms.

Lessons LearnedFrequent Questions

☐ What is the relationship between the number of algorithms and datasets/problems to do a correct statistical analysis?

☐ In multiple comparisons, the number of problems (data-sets) must be greater than the double of algorithms. With lesser data-sets, it is highly probable to not reject any null hyphotesis.

Frequent Questions

☐ Is there a maximum number of datasets/problems to be used?

There not exists a theoretical threshold, although if the number of problems is very high in relation with the number of algorithms, the results trend to be inaccurate by the central limit theorem. For pairwise comparisons, such Wilcoxon's, a maximum of 30 problems is suggested. In multiple comparisons with a control, we should indicate as a rule of thumb that $n > 8 \cdot k$ could be excessive and results in no significant comparisons.

Lessons Learned

Frequent Questions

☐ The Wilcoxon test applied several times works better than a multiple comparison test such as Holm, Is it correct to be used in these cases?

☐ The Wilcoxon test can be applied according a multiple comparison scheme, but the results obtained cannot be considered into a family which control the FWER. Each time a new comparison is conducted, the level of significance established a priori can be overcome. For this reason, the multiple comparison tests exist.

Lessons Learned Frequent Questions

☐ Can we use only the rankings obtained to justify the results?

□ With the rankings values obtained by Friedman and derivatives we can establish a clear order in the algorithms and even to measure the differences among them. However, it cannot be concluded that one proposal is better than other until the hypothesis of comparison associated to them is rejected.

Lessons Learned Frequent Questions

☐ Is it necessary to check the rejection of the null hypothesis of Friedman and derivatives before conducting a post-hoc analysis?

☐ It should be done, although by definition, it can be computed independently.

Lessons Learned

Frequent Questions

☐ When the Friedman Aligned and Quade tests are recommendable instead of classical Friedman?

The difference of power among the three methods are small and very dependent of the sample of results to be analyzed. Theoretical studies demonstrate that the Aligned Friedman and the Quade tests have better performance when we compare not more than 4 algorithms. The Quade test also assumes some risk because it considers that the more relevant problems are also those which present higher differences in performance among the methods, and it is not always true.

Lessons Learned

Frequent Questions

☐ Which post-hoc procedures should be used?

We consider that the Holm test must appear in a comparison, wheres Bonferroni does not. Hochberg and Li tests could act as a complement when their use allow us to reject more hypotheses than Holm's. All rejected hyphotesis by any procedure is correctly rejected because all procedures perform a strong control of the FWER. However, some tests, such as Li, are influenciated by the unadjusted p-values of the initial hypotheses and when the are lesser than 0.5, is the only case in which the test achieves its best performance of power.

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Statistical Analysis of Experiments in Data Mining and Computational Intelligence

OUTLINE (II)

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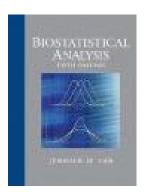
P1: S. García, <u>F. Herrera</u>, An Extension on "Statistical Comparisons of Classifiers over Multiple Data Sets" for all Pairwise Comparisons. *Journal of Machine Learning Research* 9 (2008) 2677-2694

P2: <u>J. Luengo</u>, S. García, <u>F. Herrera</u>, A Study on the Use of Statistical Tests for Experimentation with Neural Networks: Analysis of Parametric Test Conditions and Non-Parametric Tests. *Expert Systems with Applications 36 (2009) 7798-7808 doi:10.1016/j.eswa.2008.11.041*.

P3: S. García, A. Fernández, <u>J. Luengo</u>, <u>F. Herrera</u>, A Study of Statistical Techniques and Performance Measures for Genetics-Based Machine Learning: Accuracy and Interpretability. *Soft Computing 13:10 (2009) 959-977*, <u>doi:10.1007/s00500-008-0392-y</u>.

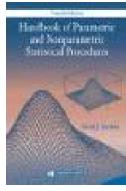
P4: S. García, <u>D. Molina</u>, <u>M. Lozano</u>, <u>F. Herrera</u>, A Study on the Use of Non-Parametric Tests for Analyzing the Evolutionary Algorithms' Behaviour: A Case Study on the CEC'2005 Special Session on Real Parameter Optimization. *Journal of Heuristics*, 15 (2009) 617-644. <u>doi: 10.1007/s10732-008-9080-4</u>.

P5: S. García, A. Fernández, <u>J. Luengo</u>, <u>F. Herrera</u>, <u>Advanced nonparametric tests</u> for multiple comparisons in the design of experiments in computational intelligence and data mining: Experimental Analysis of Power. *Information Sciences* 180 (2010) 2044–2064. <u>doi:10.1016/j.ins.2009.12.010</u>.



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Statistical Analysis of Experiments in Data Mining and Computational Intelligence

OUTLINE (II)

- Advanced non-parametric tests and case studies:
 - **■** For Multiple Comparisons involving control method
 - Post-hoc Procedures
 - Adjusted p-values
 - Detecting all pairwise differences in a multiple comparison
- Lessons Learned
 - Considerations on the use of nonparametric tests
 - **■** Recommendations on the use of nonparametric tests
 - **■** Frequent Questions
- Books of Interest and References
- Software

Software

Software for conducting multiple comparisons tests with a control



http://sci2s.ugr.es/sicidm/controlTest.zip

Read data of results of k algorithms over N case problems in CSV format. The data can correspond to accuracy, AUC or any other performance measure.

Compute the rankings through the Friedman Aligned Ranks and Quade procedures of k algorithms over N case problems.

Compute the Friedman and Iman-Davenport, Friedman Aligned-Ranks and Quade Statistics corresponding to the input data.

Show the tables with the set of hypotheses, unadjusted p-values for each comparison and adjusted level of significance for Bonferroni-Dunn, Holm and Hochberg, Hommel, Holland, Rom, Finner and Li procedures: 1 x n comparison.

Show the table with adjusted p-values for the procedures 1 x n mentioned in the previous item.

Give a report detailing the rejected hypotheses considering the levels of significance $\alpha = 0.05$ and $\alpha = 0.10$.

Software

Software for conducting all pairwise comparisons

http://sci2s.ugr.es/sicidm/multipleTest.zip



We offer a software developed in JAVA which calculates all the multiple comparisons procedures described in this talk and the JMLR paper.

It allows as input files in CSV format and obtains as output a LaTeX file with tabulated information about Friedman, Iman-Davenpor. Bonferroni-Dunn, Holm, Hochberg, Shaffer and Bergamnn-Hommel tests. It also computes and shows the adjusted p-values.

Statistical Analysis of Experiments in Data Mining and Computational Intelligence

How must I conduct statistical comparisons in my Experimental Study? On the use of Nonparametric Tests and Case Studies.

